# **Bayesian Nonparametrics**

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9.520 Class 11

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C. Frogner Bayesian Nonparametrics

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- Last time Bayesian formulation of RLS, for regression. (Basically, a normal distribution.)
- This time a more complicated probability model: the Dirichlet Process.
  - And its application to clustering.
- And also more Bayesian terminology.

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## Plan

#### • Dirichlet distribution + other basics

- The Dirichlet process
  - Abstract definition
  - Stick Breaking
  - Chinese restaurant process
- Clustering
  - Dirichlet process mixture model
  - Hierarchical Dirichlet process mixture model

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The Gamma function

$$\Gamma(z)=\int_0^\infty x^{z-1}e^{-x}dx.$$

Extends factorial function to  $\mathbb{R}^+$ :  $\Gamma(z + 1) = z\Gamma(z)$ .

#### **Beta Distribution**

$$P(x|\alpha,\beta) = \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{(\alpha-1)} (1-x)^{(\beta-1)}$$

for  $x \in [0, 1]$ ,  $\alpha > 0$ ,  $\beta > 0$ . (Mean:  $\frac{\alpha}{\alpha+\beta}$ , variance:  $\frac{\alpha\beta}{(\alpha+\beta)^2(\alpha+\beta+1)}$ .)

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# **Beta Distribution**



For large parameters the distribution is unimodal. For small parameters it favors biased binomial distributions.

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Generalizes Beta distribution to the K-dimensional simplex  $\mathbb{S}^{K}$ .

$$\mathbb{S}^{K} = \{ x \in \mathbb{R}^{K} : \sum_{i=1}^{K} x_{i} = 1, \ x_{i} \ge 0 \ \forall i \}$$

#### **Dirichlet distribution**

$$P(x|\alpha) = P(x_1, \dots, x_K) = \frac{\Gamma(\sum_{i=1}^K \alpha_i)}{\sum_{i=1}^K \Gamma(\alpha_i)} \prod_{i=1}^K (x_i)^{\alpha_i - 1}$$
  
where  $\alpha = (\alpha_1, \dots, \alpha_K), \alpha_i > 0 \ \forall i, x \in \mathbb{S}^K$ .

We write  $x \sim \text{Dir}(\alpha)$ , i.e.  $x_1, \ldots, x_K \sim \text{Dir}(\alpha_1, \ldots, \alpha_K)$ .

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# **Dirichlet Distribution**



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# Properties of the Dirichlet Distribution

$$\mathbf{E}[\mathbf{x}_i] = \frac{\alpha_i}{\sum_{j=1}^K \alpha_j}$$

Variance

Mean

$$\mathbf{Var}[x_i] = \frac{\alpha_i(\sum_{i \neq j} \alpha_j)}{(\sum_{j=1}^{K} \alpha_j)^2 (1 + \sum_{j=1}^{K} \alpha_j)}.$$

Covariance

$$\mathbf{Cov}(x_i, x_j) = \frac{\alpha_i \alpha_j}{(\sum_{j=1}^K \alpha_j)^2 (1 + \sum_{j=1}^K \alpha_j)}.$$

- Marginals:  $x_i \sim \text{Beta}(\alpha_i, \sum_{j \neq i} \alpha_j)$
- Aggregation:  $(x_1 + x_2, \dots, x_k) \sim \text{Dir}(\alpha_1 + \alpha_2, \dots, \alpha_K)$

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If you throw n balls into k bins, the distribution of balls into bins is given by the multinomial distribution.

#### Multinomial distribution

Let  $p = (p_1, ..., p_K)$  be probabilities over K categories and  $C = (C_1, ..., C_K)$  be category counts.  $C_i$  is the number of samples in the *i*th category, from n independent draws of a categorical variable with category probabilities p. Then

$$P(C|n,p) = \frac{n!}{\prod_{i=1}^{K} C_i!} \prod_{i=1}^{K} p_i^{C_i}.$$

For K = 2 this is the binomial distribution.

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# Treat the Dirichlet distribution as a distribution on probabilities: each sample $\theta \sim \text{Dir}(\alpha)$ defines a *K*-dimensional multinomial distribution.

 $\mathbf{x} \sim \mathsf{Mult}(\theta), \theta \sim \mathsf{Dir}(\alpha)$ 



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Treat the Dirichlet distribution as a distribution on probabilities: each sample  $\theta \sim \text{Dir}(\alpha)$  defines a *K*-dimensional multinomial distribution.

 $x \sim \mathsf{Mult}(\theta), \theta \sim \mathsf{Dir}(\alpha)$ 

Posterior on  $\theta$ :

 $\theta | \mathbf{x} \sim \mathsf{Dir}(\alpha + \mathbf{x})$ 

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#### Say $x \sim F(\theta)$ (the likelihood) and $\theta \sim G(\alpha)$ (the prior).

#### Conjugate prior

*G* is a **conjugate prior** for *F* if the **posterior**  $P(\theta|x, \alpha)$  is in the same family as *G*. (E.g. if *F* is Gaussian then  $P(\theta|x, \alpha)$  should also be Gaussian.)

So the Dirichlet distribution is a conjugate prior for the multinomial.

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## Plan

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  - Stick Breaking
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  - Dirichlet process mixture model
  - Hierarchical Dirichlet process mixture model

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- Parametric: fix parameters independent of data.
- **Nonparametric**: effective number of parameters can grow with the data.
- E.g. density estimation: fitting Gaussian vs. parzen windows.
- E.g. Kernel methods are nonparametric.

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Want: distribution on all K-dimensional simplices (for all K).

#### Informal Description

*X* is a space, *F* is a probability distribution on *X* and  $\mathcal{F}(X)$  is the set of all possible distributions on *X*. A **Dirichlet Process** gives a distribution over  $\mathcal{F}(X)$ . A sample

path from a DP is an element  $F \in \mathcal{F}(X)$ . *F* can be seen as a (random) probability distribution on *X*.

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Want: distribution on all K-dimensional simplices (for all K).

#### Formal Definition

Let *X* be a space and *H* be the base measure on *X*. *F* is a sample from the Dirichlet Process  $DP(\alpha, H)$  on *X* if its finite-dimensional marginals have the Dirichlet distribution:

$$(F(B_1),\ldots,F(B_K)) \sim \mathsf{Dir}(\alpha H(B_1),\ldots,\alpha H(B_2))$$

for all partitions  $B_1, \ldots, B_K$  of X (for any K).

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Explicit construction of a DP.

Let  $\alpha > 0$ ,  $(\pi_i)_{i=1}^{\infty}$  such that

$$p_i = \beta_i \prod_{j=1}^{i-1} (1 - \beta_j) = \beta_i (1 - \sum_{j=1}^{i-1} p_j)$$

where  $\beta_i \sim Beta(1, \alpha)$ , for all *i*. Let *H* be a distribution on *X* and define

$$F = \sum_{i=1}^{\infty} p_i \delta_{\theta_i}$$

where  $\theta_i \sim H$ , for all *i*.

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# Stick Breaking Construction: Interpretation



The weights  $\pi$  partition a unit-length *stick* in an infinite set: the *i*-th weight is a random proportion  $\beta_i$  of the stick remaining after sampling the first *i* – 1 weights.

It is possible to prove (Sethuraman '94) that the previous construction returns a DP and conversely a Dirichlet process is discrete almost surely.

There is an infinite (countable) set of tables.

- First customer sits at the first table.
- Customer *i* sits at table *j* with probability

$$\frac{n_j}{\alpha+i+1},$$

where  $n_j$  is the number of customers at table *j*, and *i* sits at the first open table with probability

$$\frac{\alpha}{\alpha + i + 1}$$

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Note that  $\mathbb{E}[\beta_i] = 1/(1 + \alpha)$ .

- for small  $\alpha$ , the first few components will have all the mass.
- for large α, F approaches the distribution H assigning uniform weights to the samples θ<sub>i</sub>.

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# It is possible to prove (Antoniak '77??) that the number of components with positive count grows as

 $\alpha \log n$ 

as we increase the number of samples *n*.

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# Clustering with the *K*-dimensional Dirichlet: take each sample $\theta \sim \text{Dir}(\alpha)$ to define a *K*-dimensional categorical (instead of multinomial) distribution.

#### $x \sim G(\phi), \phi \sim \operatorname{Cat}(\theta), \theta \sim \operatorname{Dir}(\alpha)$

(*G* is a a distribution on observation space *X*, say, Gaussian.)  $\theta_i$  is the probability of *x* coming from the *i*th cluster.

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Clustering with the *K*-dimensional Dirichlet: take each sample  $\theta \sim \text{Dir}(\alpha)$  to define a *K*-dimensional categorical (instead of multinomial) distribution.

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Clustering with the Dirichlet Process: take each sample  $\theta \sim DP(\alpha, H)$  to define a *K*-dimensional categorical (instead of multinomial) distribution.

$$x \sim G(\phi), \phi \sim Cat(\theta), \theta \sim DP(\alpha, H)$$

(*G* is a distribution on observation space *X*, say, Gaussian. *H* can be uniform on  $\{1, \ldots, K\}$ .)

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Clustering with the Dirichlet Process:

$$x \sim G(\phi), \phi \sim Cat(\theta), \theta \sim \mathsf{DP}(\alpha, H)$$

This is the Dirichlet Process mixture model.

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# What if we want to model **grouped data**, each group corresponding to a different DP mixture model?

#### Hierarchical Dirichlet Process

For each  $i \in \{1, \ldots, n\}$ , draw  $x_i$  according to

 $x_i \sim G(\phi), \phi \sim \operatorname{Cat}(\theta), \theta \sim \operatorname{DP}(\alpha, H_0), \alpha \sim \operatorname{DP}(\gamma, H).$ 

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#### **Hierarchical Dirichlet Process**

For each  $i \in \{1, \ldots, n\}$ , draw  $x_i$  according to

 $x_i \sim G(\phi), \phi \sim Cat(\theta), \theta \sim DP(\alpha, H_0), \alpha \sim DP(\gamma, H).$ 

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- Dirichlet distribution gives a distribution over the *K*-simplex.
- Dirichlet is conjugate to the multinomial, which makes inference in the Dirichlet/multinomial model easy.
- Dirichlet process generalizes the Dirichlet distribution to countably infinitely many components.
  - Every finite marginal of the DP is Dirichlet distributed.
- Complexity of the DP is controlled by the strength parameter  $\alpha$ .
- The posterior distribution cannot be found analytically. Approximate inference is needed.

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This lecture heavily draws (sometimes literally) from the list of references below, which we suggest as further readings.

Figures are taken either from Sudderth PhD thesis or Teh Tutorial.

Main references/sources:

- Yee Whye Teh, *Tutorial in the Machine Learning Summer School*, and his notes *Dirichlet Processes*.
- Erik Sudderth, PhD Thesis.
- Gosh and Ramamoorthi, *Bayesian Nonparametrics*, (book).

See also:

- Soubin Ghahramani, Tutorial ICML.
- Michael Jordan, Nips Tutorial.
- Rasmussen, Williams, Gaussian Processes for Machine Learning, (book).
- Ferguson, paper in Annals of Statistics.
- Sethuraman, paper in *Statistica Sinica*.
- Berlinet, Thomas-Agnan, *RKHS in Probability and Statistics*, (book).

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#### APPENDIX



A partition of X is a collection of subsets  $B_1, \ldots, B_N$  is such that, if  $B_i \cap B_j = \emptyset$ ,  $\forall i \neq j$  and  $\bigcup_{i=1}^N B_i = X$ .

#### Definition (Existence Theorem)

Let  $\alpha > 0$  and *H* a probability distribution on *X*. One can prove that there exists a unique distribution  $DP(\alpha, H)$ on  $\mathcal{F}(X)$  such that, if  $F \sim DP(\alpha, H)$  and  $B_1, \ldots, B_N$  is a partition of *X* then

$$(F(B_1),\ldots,F(B_N)) \sim \text{Dir}(\alpha H(B_1),\ldots,\alpha H(B_N)).$$

The above result is proved (Ferguson '73) using Kolmogorov's Consistency theorem (Kolmogorov '33).

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# **Dirichlet Processes Illustrated**



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Hereafter  $F \sim DP(\alpha, H)$  and A is a measurable set in X.

• Expectation:  $\mathbb{E}[F(A)] = \alpha H(A)$ .

• Variance: 
$$\mathbb{V}[F(A)] = \frac{H(A)(1-H(A))}{\alpha+1}$$

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## Properties of Dirichlet Processes (cont.)

Posterior and Conjugacy: let x ~ F and consider a fixed partition B<sub>1</sub>,..., B<sub>N</sub>, then

$$P(F(B_1),\ldots,F(B_N)|x \in B_k) =$$
  
Dir( $\alpha H(B_1),\ldots,\alpha H(B_k) + 1,\ldots,\alpha H(B_N)$ ).

It is possible to prove that if  $S = (x_1, ..., x_n) \sim F$ , and  $F \sim DP(\alpha, H)$ , then

$$P(F|S,\alpha,H) = DP\left(\alpha + n, \frac{1}{n+\alpha}\left(\alpha H + \sum_{i=1}^{n} \delta_{x_i}\right)\right)$$

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# A Qualitative Reasoning

From the form of the posterior we have that

$$\mathbb{E}(F(A)|S,\alpha,H) = \frac{1}{n+\alpha} \left( \alpha H(A) + \sum_{i=1}^{n} \delta_{x_i}(A) \right).$$

If  $\alpha < \infty$  and  $n \to \infty$  one can argue that

$$\mathbb{E}(F(A)|S,\alpha,H) = \sum_{i=1}^{\infty} \pi_i \delta_{x_i}(A)$$

where  $(\pi_i)_{i=1}^{\infty}$  is the sequence corresponding to the limit  $\lim_{n\to\infty} C_i/n$  of the empirical frequencies of the observations  $(x_i)_{i=1}^{\infty}$ .

If the posterior concentrates about its mean the above reasoning suggests that the obtained distribution is discrete.