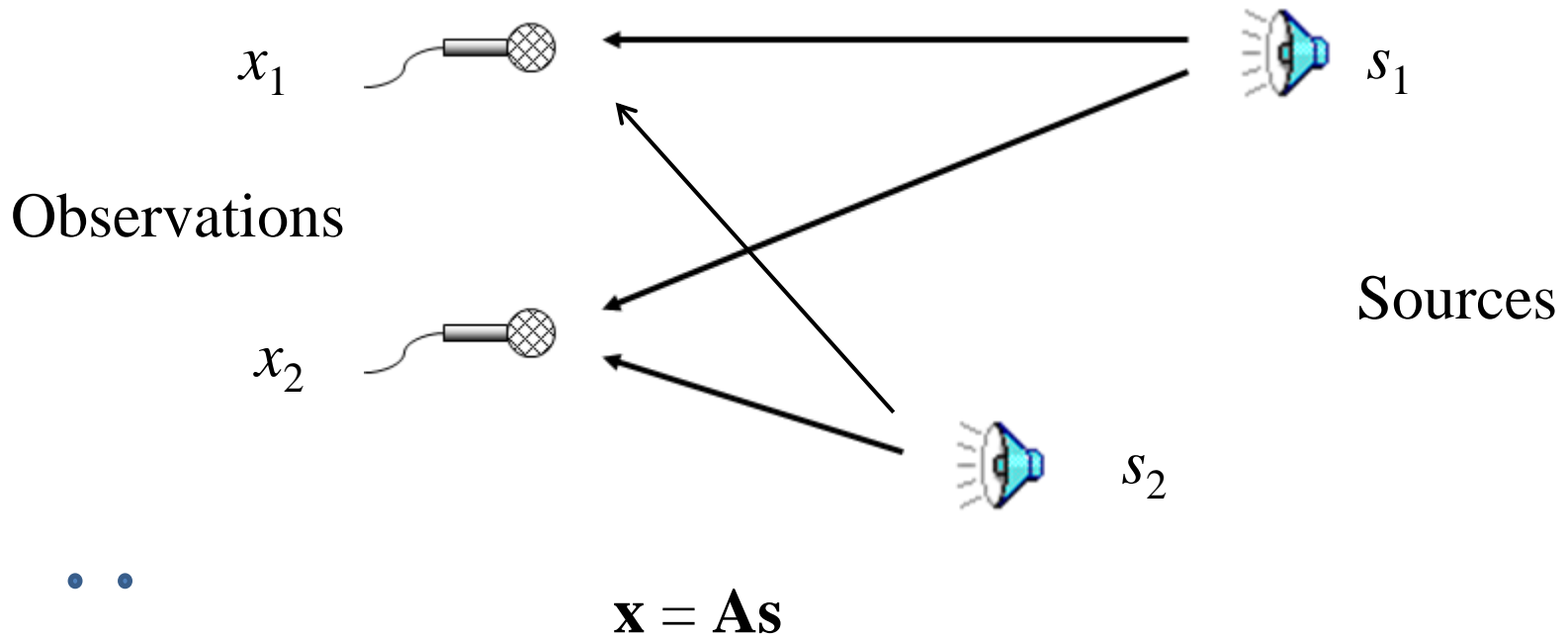


***JCA***

# ICA: 2-D examples



$$\begin{matrix} \text{---} \\ \text{---} \end{matrix} = \begin{matrix} \text{---} \\ \text{---} \end{matrix} \times \begin{matrix} \text{---} \\ \text{---} \end{matrix}$$

$X_{2 \times n}$        $A_{2 \times 2}$        $S_{2 \times n}$

# Independent Components Analysis

$$X_1 = a_{11}S_1 + a_{12}S_2 + \dots + a_{1p}S_p$$

$$X_2 = a_{21}S_1 + a_{22}S_2 + \dots + a_{2p}S_p$$

⋮

$$X_p = a_{p1}S_1 + a_{p2}S_2 + \dots + a_{pp}S_p$$

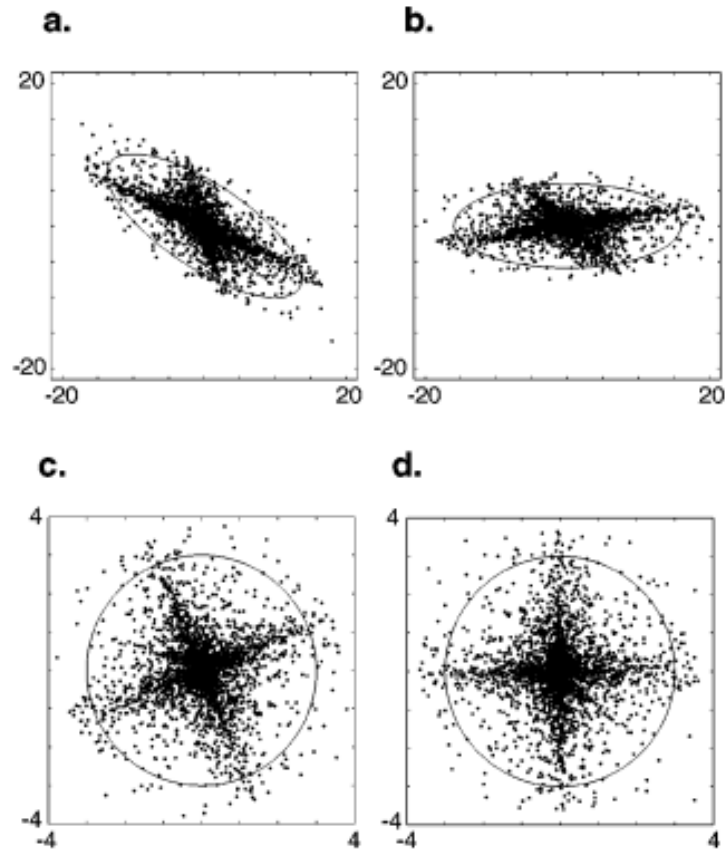
$$X = AS$$

If we knew A we could solve for the sources S

But we have to solve for *both*

We will look for a solution that will make S *independent*

# PCA and ICA



$$\mathbf{X} = \mathbf{A}\mathbf{S}$$

- Getting a simpler form
- We can always express  $\mathbf{A}$  by SVD as  $\mathbf{U}\mathbf{\Sigma}\mathbf{V}^T$
- $\mathbf{U}$  and  $\mathbf{V}$  are orthonormal and  $\mathbf{\Sigma}$  is diagonal
- (we don't know any of them)
- So now  $\mathbf{X} = \mathbf{U}\mathbf{\Sigma}\mathbf{V}^T \mathbf{S}$
  
- Taking the covariance matrix of the data:
- $\mathbf{X}\mathbf{X}^T = \mathbf{U}\mathbf{\Sigma}\mathbf{V}^T \mathbf{S} \mathbf{S}^T \mathbf{V}\mathbf{\Sigma}\mathbf{U}^T$
- We can assume that  $\mathbf{S}\mathbf{S}^T = \mathbf{I}$
- They are independent, therefore uncorrelated.
- We can assume all of length = 1
- This is just scaling; we can scale  $\mathbf{S}$  and  $\mathbf{A}$

- $X = AS$
- $A = U\Sigma V^T$  (the SVD of A)
- $X = U\Sigma V^T S$
- $XX^T = U\Sigma V^T S S^T V \Sigma U^T$  with  $SS^T = I$
- $XX^T = U\Sigma^2 U$

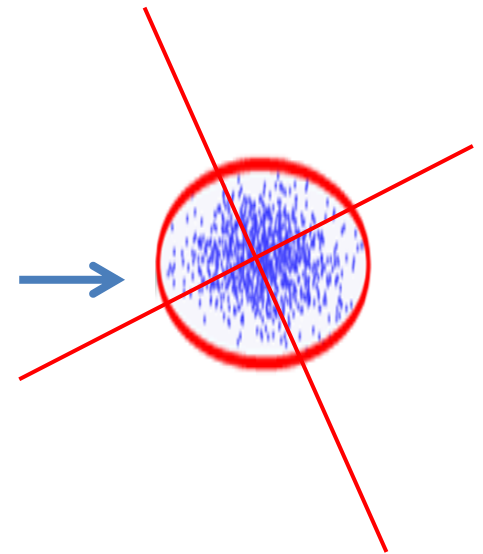
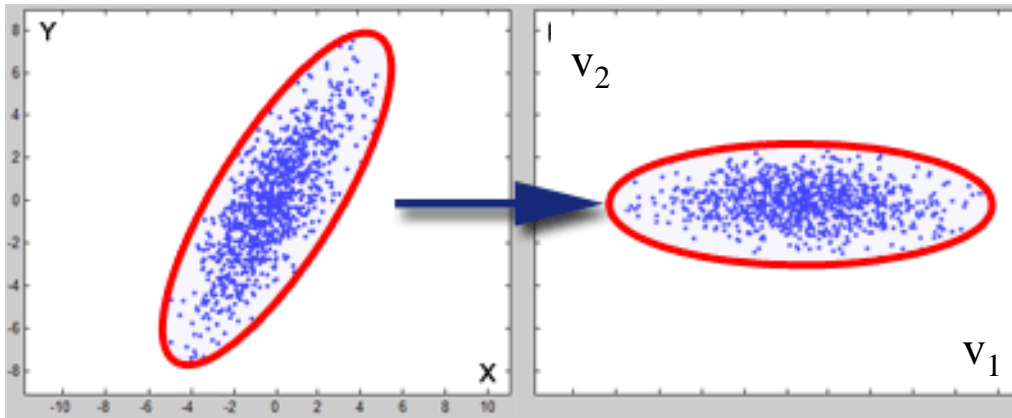
With the *same*  $U, \Sigma$  we used for A above

- $XX^T$  is known, so we can find the  $U, \Sigma$  of A from the data
- (by diagonalizing  $XX^T = U \Lambda U^T$  )

# ICA procedure

- Looking for  $X = AS$  with  $S$  independent
- Start by whitening  $X$ :
- Do PCA, then:  $X' \leftarrow \Sigma^{-1}U^T X$
- In the new data solve for  $X' = VS$
- Both  $V, S$  unknown, but  $V$  is rotation, and  $S$  are independent.
- Search over rotations and test for independence
- For a given  $V$ ,  $S$  is easy to obtain, we need some measure of independence

# Whitening the data



Perform PCA

Re-scale the coordinates by their variance

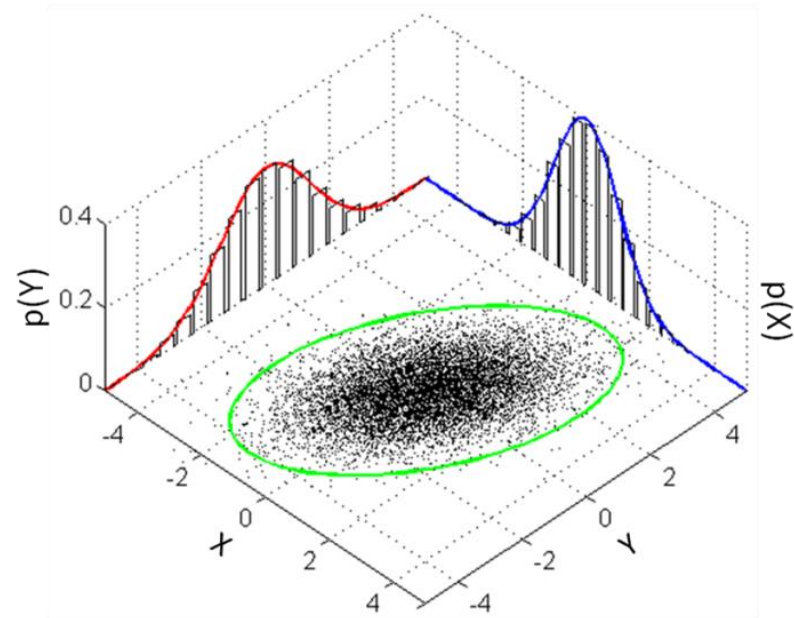
ICA: Final step – look for rotation that will make  $S$  as independent as possible



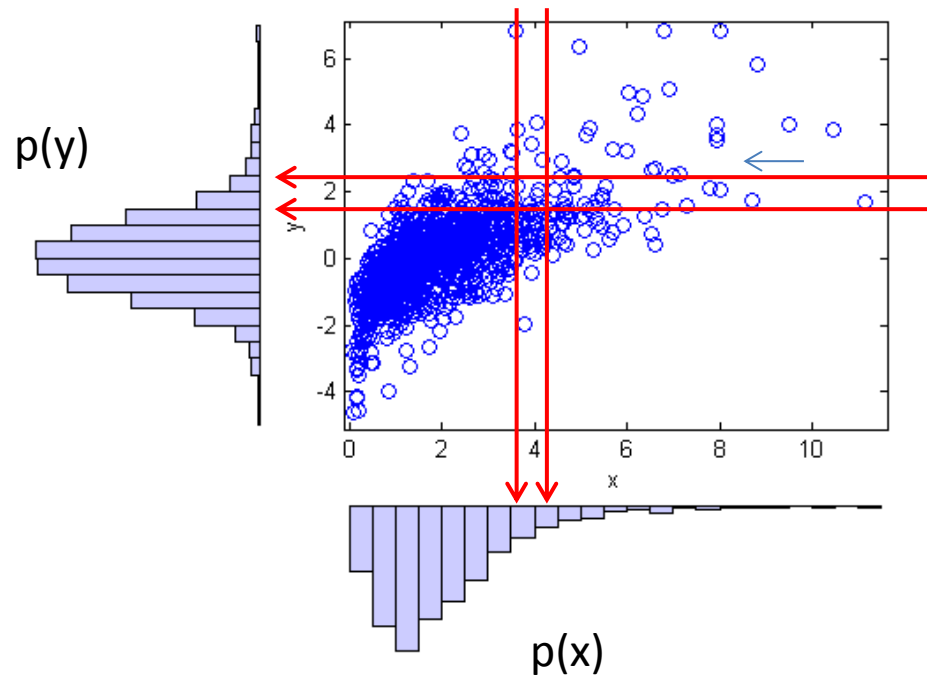
# Testing for Independence

- Suppose that a source produces variables  $(x_1 y_1) (x_2 y_2) \dots$ .
- It is straightforward to test if they are correlated or not by  $\sum x_i y_i = 0$
- In practice,  $\sum x_i y_i > \epsilon$
- How to test independence?
  
- Several methods, describe briefly one.

# 1-D projection



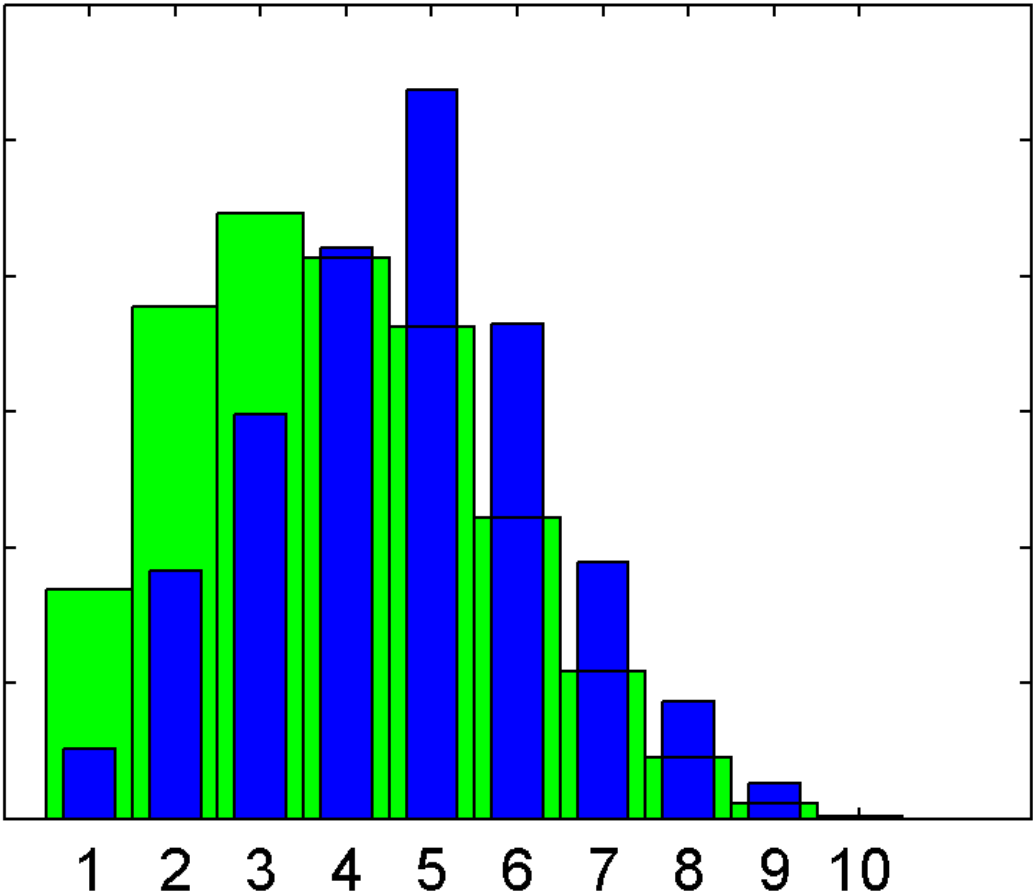
# Testing independence



$$p(x,y) = p(x) p(y)$$

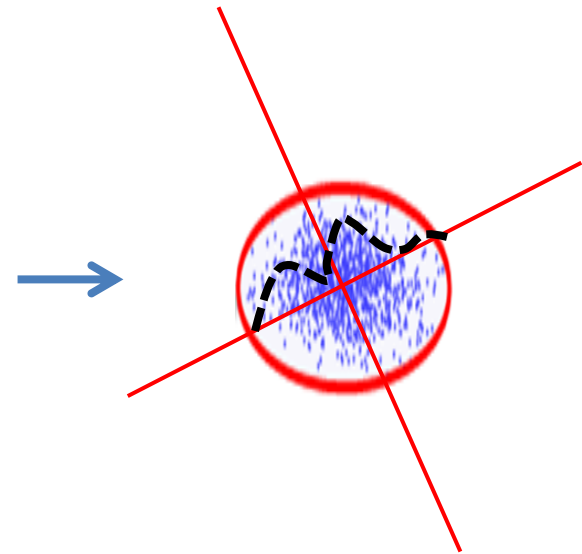
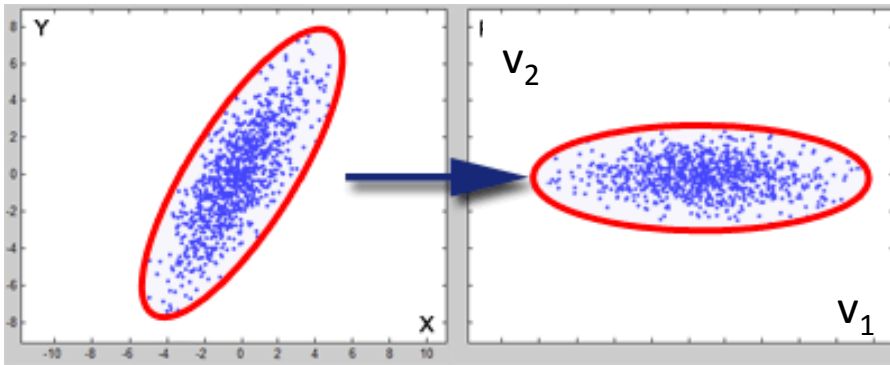
- In principle for each pair  $x_i y_j$  verify that  $p(x_i y_j) = p(x_i) p(y_i)$
- We have many pairs, how to use them together in an efficient test
- We look at the two distributions  $p(x,y)$  and  $q(x,y) = p(x)p(y)$
- We want to test if they the same (or very close)
  
- How to compare two distributions?

Two distributions – how different are they?



# Testing for Independence

- Use the KL divergence: Kullback-Leibler
- $KL(p||q) = \sum [ p \log (p/q)]$
- Non-negative, it is 0 only iff they are the same.
  
- In our case
- $KL [p(x, y) || p(x) p(y)] = \sum [p(x, y) \log (p(x, y)/p(x) p(y))] =$
- $\sum p(x, y) \log p(x, y) - (\sum p(x, y) \log p(x) + \sum p(x, y) \log p(y))$
- $= -H(p(x, y)) + [H(p(x)) + H(p(y))]$
- 
- $\sum H_i - H$
  
- $H$  is constant, minimize  $\sum H_i$  (marginal distribution after rotation)



Final step: optimize iteratively over rotation. For each rotation project the data on the axes and measure  $H_i$  of the projections.

## Technical difficulties:

- Minimizing  $\sum H_i$  on all the axes
- Non-convex, complex, minimization
- Estimating entropy  $H$ , requires enough samples, sensitive to outliers
- Various algorithms to optimize the numeric process
- FastICA (Hyvärinen), Proceeds one component at a time, then combines them



# Equivalent Criterion

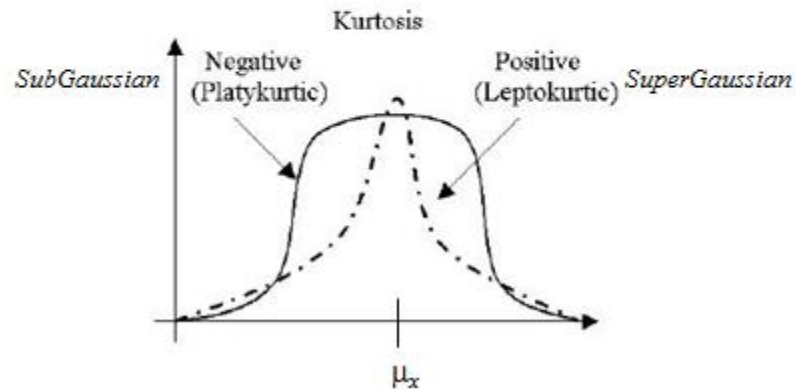
- Rotation that maximizes  $H - \sum H_i$  also maximizes the “non-Gaussianity” of the transformed data.
- 
- Non-Gaussianity (‘negentropy’): as the Kullback-Leibler divergence of a distribution from a Gaussian distribution with equal variance.
- 
- Non-gaussianity is also measured by Kurtosis
- 
- Family of algorithms that maximize Kurtosis rather than marginal entropies

# Kurtosis

Higher order moments (4<sup>th</sup>-kurtosis)

$$\hat{\kappa}(\mathbf{x}) = \frac{1}{M} \sum_{i=1}^M \left[ \frac{x_i - \hat{\mu}_x}{\hat{\sigma}} \right]^4$$

Gaussians are *mesokurtic* with  $\kappa = 3$



Non-Gaussianity: Kurtosis should be far from 3

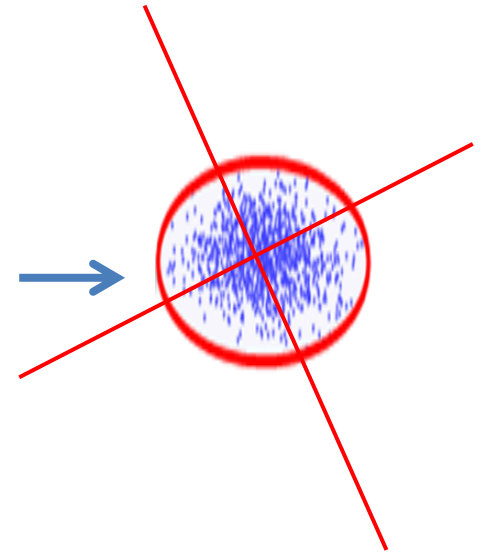
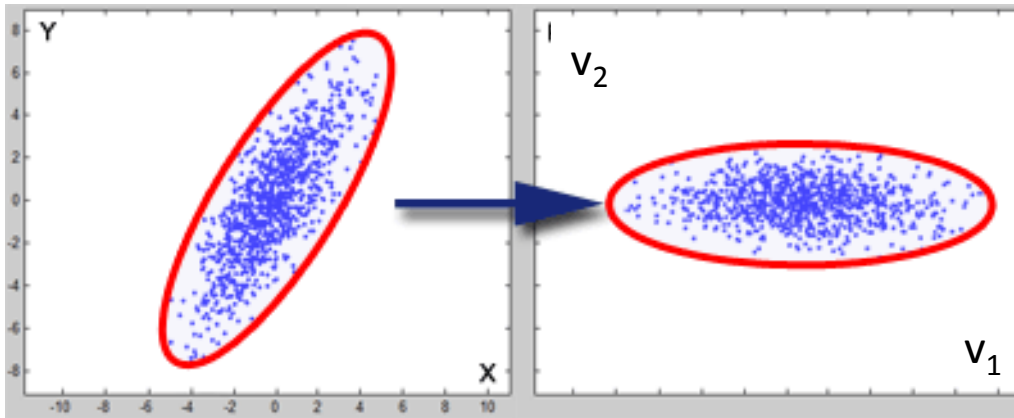
A family of algorithms that use Kurtosis rather than marginal entropies

# On Whitening the Data

- An important step in general, additional comments:
- The data matrix  $XX^T$  can be expressed as:  $U\Lambda U^T$
- 
- Whitening  $X$  is:
- $X_W = \Lambda^{-1/2} U^T X$
- 
- We can check:
- $X_W X_W^T = \Lambda^{-1/2} U^T X X^T U \Lambda^{-1/2}$
- 
- Substituting  $XX^T$
- 
- $\Lambda^{-1/2} U^T U \Lambda U^T U \Lambda^{-1/2} = I$

# On Whitening the Data

- Whitening:  $\mathbf{X}_W = \Lambda^{-1/2} \mathbf{U}^T \mathbf{X}$
- *Regularization:*
- $\Lambda^{-1/2}$  is a diagonal matrix with  $1/(\text{sqrt } \lambda_i)$  on the diagonal
- This is regularized to  $1/(\text{sqrt } \lambda_i + \epsilon)$
- *ZCA (zero-phase whitening)*
- 
- Whitening is non-unique.
- Any rotation will leave it whitened (next slide)
- 
- Taking in particular  $\mathbf{U}$  from the data matrix:
- 
- $\mathbf{X}_{ZCA} = \mathbf{U} \Lambda^{-1/2} \mathbf{U}^T \mathbf{X}$
- 
- From all whitened  $\mathbf{X}_W$ , this is the closest to the original  $\mathbf{X}$ .



After whitening, added rotation leaves the data whitened

Next: Performing the ICA on image patches:

- **The “independent components” of natural scenes are edge filters**
- Bell and Sejnowski Vision Research 1997