MIT 9.520/6.860, Fall 2019 Statistical Learning Theory and Applications

Class 03: Regularized Least Squares

Lorenzo Rosasco

Learning problem and algorithms

Solve

$$\min_{f \in \mathcal{F}} L(f), \qquad L(f) = \mathbb{E}_{(x,y) \sim P}[\ell(y,f(x))],$$

given only

$$S_n = (x_1, y_1), \dots, (x_n, y_n) \sim P^n.$$

Learning algorithm

$$S_n \to \widehat{f} = \widehat{f}_{S_n}$$
 ,

 \widehat{f} estimates f_P given the observed examples S_n .

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How can we design a learning algorithm?

Algorithm design: complexity and regularization

The design of most algorithms proceed as follows:

 \triangleright Pick a (possibly large) class of function \mathcal{H} , ideally

$$\min_{f \in \mathcal{H}} L(f) = \min_{f \in \mathcal{F}} L(f).$$

▶ Define a procedure $A_{\gamma}(S_n) = \hat{f}_{\gamma} \in \mathcal{H}$ to explore the space \mathcal{H} .

Empirical risk minimization

A classical example (called M-estimation in statistics).

Consider $(\mathcal{H}_{\gamma})_{\gamma}$ such that

$$\mathcal{H}_1 \subset \mathcal{H}_2, \dots \mathcal{H}_{\gamma} \subset \dots \mathcal{H}.$$

Then, let

$$\widehat{f}_{\gamma} = \min_{f \in \mathcal{H}_{\gamma}} \widehat{L}(f), \qquad \qquad \widehat{L}(f) = \frac{1}{n} \sum_{i=1}^{n} \ell(y_i, f(x_i)).$$

This is the idea we discuss next.

Linear functions

Let \mathcal{H} be the space of linear functions

$$f(x) = w^{T}x$$
.

Then,

- ightharpoonup f \leftrightarrow w is one to one,
- inner product $\langle f, \overline{f} \rangle_{\mathcal{H}} := \mathbf{w}^{\top} \overline{\mathbf{w}},$
- $\qquad \qquad \operatorname{norm/metric} \ \left\| f \overline{f} \right\|_{\mathcal{H}} := \| w \overline{w} \|.$

Linear functions are the conceptual building block of most functions.

Linear least squares

ERM with least squares also called ordinary least squares (OLS)

$$\min_{w \in \mathbb{R}^d} \frac{1}{n} \underbrace{\sum_{i=1}^n (y_i - w^\top x_i)^2}_{\widehat{L}(w)}.$$

- ► Statistics later...
- ...now computations.

Matrices and linear systems

Let $\widehat{X} \in \mathbb{R}^{nd}$ and $\widehat{Y} \in \mathbb{R}^{n}$. Then

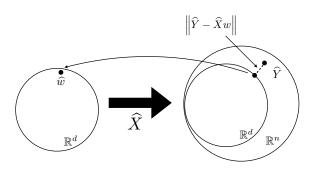
$$\frac{1}{n}\sum_{i=1}^n(y_i-w^\top x_i)^2=\frac{1}{n}\left\|\widehat{Y}-\widehat{X}w\right\|^2.$$

This is the least squares problem associated to the linear system

$$\widehat{X}w=\widehat{Y}.$$

Overdetermined lin. syst.

n > d



$$\nexists \widehat{w} \quad s.t. \quad \widehat{X}w = \widehat{Y}$$

Least squares solutions

From the optimality conditions

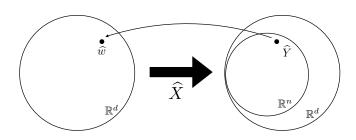
$$\nabla_{\mathbf{w}} \frac{1}{\mathbf{n}} \left\| \widehat{\mathbf{Y}} - \widehat{\mathbf{X}} \mathbf{w} \right\|^2 = 0$$

we can derive the normal equation

$$\widehat{X}^{\top}\widehat{X}w = \widehat{X}^{\top}\widehat{Y} \qquad \Leftrightarrow \qquad \widehat{w} = (\widehat{X}^{\top}\widehat{X})^{-1}\widehat{X}^{\top}\widehat{Y}.$$

Underdetermined lin. syst.

n < d



$$\exists \ \widehat{\mathbf{w}} \quad \text{s.t.} \quad \widehat{\mathbf{X}} \mathbf{w} = \widehat{\mathbf{Y}}$$

possibly not unique...

Minimal norm solution

There can be many solutions

$$\widehat{X}\widehat{w} = \widehat{Y}, \quad \text{and} \quad \widehat{X}w_0 = 0 \quad \Rightarrow \widehat{X}\big(\widehat{w} + w_0\big) = \widehat{Y}.$$

Consider

$$\min_{w \in \mathbb{R}^d} \|w\|^2, \quad \text{subj. to} \quad \widehat{X}w = \widehat{Y}.$$

Using the method of Lagrange multipliers, the solution is

$$\widehat{w} = \widehat{X}^{\top} (\widehat{X} \widehat{X}^{\top})^{-1} \widehat{Y}.$$

Pseudoinverse

$$\widehat{w} = \widehat{X}^{\dagger} \widehat{Y}$$

For n > d, (independent columns)

$$\widehat{X}^{\dagger} = (\widehat{X}^{\top} \widehat{X})^{-1} \widehat{X}^{\top}.$$

For n < d, (independent rows)

$$\widehat{X}^{\dagger} = \widehat{X}^{\top} (\widehat{X} \widehat{X}^{\top})^{-1}.$$

Spectral view

Consider the SVD of \widehat{X}

$$\widehat{X} = USV^{\top} \quad \Leftrightarrow \quad \widehat{X}w = \sum_{j=1}^{r} s_{j}(v_{j}^{\top}w)u_{j},$$

here $r \le n \wedge d$ is the rank of \widehat{X} .

Then,

$$\widehat{\boldsymbol{w}}^{\dagger} = \widehat{\boldsymbol{X}}^{\dagger} \widehat{\boldsymbol{Y}} = \sum_{j=1}^{r} \frac{1}{s_{j}} (\boldsymbol{u}_{j}^{\top} \widehat{\boldsymbol{Y}}) \boldsymbol{v}_{j}.$$

Pseudoinverse and bias

$$\widehat{\boldsymbol{w}}^{\dagger} = \widehat{\boldsymbol{X}}^{\dagger} \widehat{\boldsymbol{Y}} = \sum_{i=1}^{r} \frac{1}{s_{j}} (\boldsymbol{u}_{j}^{\top} \widehat{\boldsymbol{Y}}) \boldsymbol{v}_{j}.$$

 $(v_j)_j$ are principal components of $\widehat{X}\!:$ OLS "likes" principal components.

Not all linear functions are the same for OLS!

The pseudoinverse introduces a bias towards certain solutions.

From OLS to ridge regression

Recall, it also holds,

$$\widehat{\mathbf{X}}^{\dagger} = \lim_{\lambda \to 0_+} (\widehat{\mathbf{X}}^{\top} \widehat{\mathbf{X}} + \lambda \mathbf{I})^{-1} \widehat{\mathbf{X}}^{\top} = \lim_{\lambda \to 0_+} \widehat{\mathbf{X}}^{\top} (\widehat{\mathbf{X}} \widehat{\mathbf{X}}^{\top} + \lambda \mathbf{I})^{-1}.$$

Consider for $\lambda > 0$,

$$\widehat{\mathbf{w}}_{\lambda} = (\widehat{\mathbf{X}}^{\top} \widehat{\mathbf{X}} + \lambda \mathbf{I})^{-1} \widehat{\mathbf{X}}^{\top} \widehat{\mathbf{Y}}.$$

This is called ridge regression.

Spectral view on ridge regression

$$\widehat{\boldsymbol{w}}_{\lambda} = (\widehat{\boldsymbol{X}}^{\top}\widehat{\boldsymbol{X}} + \lambda \boldsymbol{I})^{-1}\widehat{\boldsymbol{X}}^{\top}\widehat{\boldsymbol{Y}}$$

Considering the SVD of \widehat{X} ,

$$\widehat{\mathbf{w}}_{\lambda} = \sum_{j=1}^{r} \frac{\mathbf{s}_{j}}{\mathbf{s}_{j}^{2} + \lambda} (\mathbf{u}_{j}^{\top} \widehat{\mathbf{Y}}) \mathbf{v}_{j}.$$

Ridge regression as filtering

$$\widehat{w}_{\lambda} = \sum_{j=1}^{r} \frac{s_{j}}{s_{j}^{2} + \lambda} (u_{j}^{\top} \widehat{Y}) v_{j}$$

The function

$$F(s) = \frac{s}{s^2 + \lambda},$$

acts as a low pass filter (low frequencies= principal components).

- ▶ For s small, $F(s) \approx 1/\lambda$.
- ▶ For s big, $F(s) \approx 1/s$.

Ridge regression as ERM

$$\widehat{\boldsymbol{w}}_{\lambda} = (\widehat{\boldsymbol{X}}^{\top}\widehat{\boldsymbol{X}} + \lambda \boldsymbol{I})^{-1}\widehat{\boldsymbol{X}}^{\top}\widehat{\boldsymbol{Y}}$$

is the solution of

$$\min_{w \in \mathbb{R}^d} \underbrace{ \left\| \widehat{Y} - \widehat{X} w \right\|^2 + \lambda \|w\|^2}_{\widehat{L}_{\lambda}(w)}.$$

It follows from,

$$\Delta \widehat{L}_{\lambda}(w) = -\frac{2}{n} \widehat{X}^{\top} (\widehat{Y} - \widehat{X}w) + 2\lambda w = 2(\frac{1}{n} \widehat{X}^{\top} \widehat{X} + \lambda I)w - \frac{2}{n} \widehat{X}^{\top} \widehat{Y}.$$

Ridge regression as ERM

ERM interpretation suggests the rescaling

$$\widehat{\boldsymbol{w}}_{\lambda} = (\widehat{\boldsymbol{X}}^{\top} \widehat{\boldsymbol{X}} + {\color{red}\boldsymbol{n}} \lambda \boldsymbol{I})^{-1} \widehat{\boldsymbol{X}}^{\top} \widehat{\boldsymbol{Y}}$$

since

$$\min_{w \in \mathbb{R}^d} \frac{1}{n} \frac{\left\|\widehat{Y} - \widehat{X}w\right\|^2 + \lambda \|w\|^2}{\widehat{L}_{\lambda}(w)}.$$

Related ideas

Tikhonov

$$\min_{w \in \mathbb{R}^d} \frac{1}{n} \left\| \widehat{Y} - \widehat{X}w \right\|^2 + \lambda \|w\|^2$$

Morozov

$$\min_{\mathbf{w} \in \mathbb{R}^d} \|\mathbf{w}\|^2 \qquad \text{subj. to} \qquad \frac{1}{n} \left\| \widehat{\mathbf{Y}} - \widehat{\mathbf{X}} \mathbf{w} \right\|^2 \leq \delta$$

Ivanov

$$\min_{\mathbf{w} \in \mathbb{R}^d} \frac{1}{n} \left\| \widehat{\mathbf{Y}} - \widehat{\mathbf{X}} \mathbf{w} \right\|^2, \quad \text{subj. to} \quad \|\mathbf{w}\|^2 \le R$$

Ridge regression and SRM

The constraint

$$||\mathbf{w}||^2 \le \mathbf{R}$$

- restricts the search of solution,
- ▶ shrinks the solution coefficients.

Different views on regularization

$$\widehat{\mathbf{w}} = \widehat{\mathbf{X}}^{\dagger} \widehat{\mathbf{Y}} \qquad \qquad \widehat{\mathbf{w}}_{\lambda} = \big(\widehat{\mathbf{X}}^{\top} \widehat{\mathbf{X}} + \lambda \mathbf{I}\big)^{-1} \widehat{\mathbf{X}}^{\top} \widehat{\mathbf{Y}}$$

$$\min_{\boldsymbol{w} \in \mathbb{R}^d} \min_{\boldsymbol{s.t.} \ \widehat{\boldsymbol{X}} \boldsymbol{w} = \widehat{\boldsymbol{Y}}} \|\boldsymbol{w}\|^2 \qquad \quad \min_{\boldsymbol{w} \in \mathbb{R}^d} \frac{1}{n} \sum_{i=1}^n (y_i - \boldsymbol{w}^\top \boldsymbol{x}_i)^2 + \lambda \, \|\boldsymbol{w}\|^2$$

- ► Introduces a bias towards certain solutions: small norm/principal components,
- ightharpoonup controls the stability of the solution .

Complexity of ridge regression

Back to computations.

Solving

$$\widehat{\boldsymbol{w}}^{\lambda} = (\widehat{\boldsymbol{X}}^{\top} \widehat{\boldsymbol{X}} + \lambda \boldsymbol{I})^{-1} \widehat{\boldsymbol{X}}^{\top} \widehat{\boldsymbol{Y}}$$

requires essentially (using a direct solver)

- ightharpoonup time $O(nd^2 + d^3)$,
- ▶ memory $O(nd \lor d^2)$.

What if $n \ll d$?

Representer theorem in disguise

A simple observation

Using SVD we can see that

$$(\widehat{X}^{\top}\widehat{X} + \lambda I)^{-1}\widehat{X}^{\top} = \widehat{X}^{\top}(\widehat{X}\widehat{X}^{\top} + \lambda I)^{-1}$$

More on complexity

Then

$$\widehat{\boldsymbol{w}}_{\lambda} = \widehat{\boldsymbol{X}}^{\top} (\widehat{\boldsymbol{X}} \widehat{\boldsymbol{X}}^{\top} + \lambda \boldsymbol{I})^{-1} \widehat{\boldsymbol{Y}}.$$

requires essentially (using a direct solver)

- time $O(n^2d + n^3)$,
- ▶ memory $O(nd \lor n^2)$.

Representer theorem

Note that

$$\widehat{w}_{\lambda} = \widehat{X}^{\top} \underbrace{\left(\widehat{X}\widehat{X}^{\top} + \lambda I\right)^{-1} \widehat{Y}}_{c \in \mathbb{R}^{n}} = \sum_{i=1}^{n} x_{i} c_{i}.$$

The coefficients vector is a linear combination of the input points.

Then

$$\widehat{f}_{\lambda}(x) = x^{\top} \widehat{w}_{\lambda} = x^{\top} \widehat{X}^{\top} c = \sum_{i=1}^{n} x^{\top} x_{i} c_{i}$$

The function we obtain is a linear combination of inner products.

This will be the key to nonparametric learning.

Summing up

- ► From OLS to ridge regression
- ▶ Different views: (spectral) filtering and ERM
- ► Regularization and bias.

TBD

- ▶ Beyond linear models.
- ▶ Optimization.
- ▶ Model selection.