

The Value of Feedback for Decentralized Detection in Large Sensor Networks

Wee Peng Tay, *Member, IEEE*, John N. Tsitsiklis, *Fellow, IEEE*

Abstract—We consider the decentralized binary hypothesis testing problem in networks with feedback, where some or all of the sensors have access to compressed summaries of other sensors’ observations. We study certain two-message feedback architectures, in which every sensor sends two messages to a fusion center, with the second message based on full or partial knowledge of the first messages of the other sensors. Under either a Neyman-Pearson or a Bayesian formulation, we show that the asymptotically optimal (in the limit of a large number of sensors) detection performance (as quantified by error exponents) does not benefit from the feedback messages.

Index Terms—Decentralized detection, feedback, error exponent, sensor networks.

I. INTRODUCTION

In the problem of decentralized detection, each sensor makes an observation, and sends a summary of that observation by first applying a quantization function to its observation, and then communicating the result to a fusion center. The fusion center makes a final decision based on all the sensor messages. The goal is to design the sensor quantization functions and the fusion rule so as to minimize a cost function, such as the probability of an incorrect final decision. This problem has been widely studied for various tree architectures, including the parallel configuration [1]–[10], tandem networks [11]–[14], and bounded height tree architectures [15]–[22]. For sensor observations not conditionally independent given the hypothesis, the problem of designing the quantization functions is known to be NP-hard [23]. For this reason, most of the literature assumes that the sensor observations are conditionally independent.

Non-tree networks are harder to analyze because the different messages received by a sensor are not in general conditionally independent. While some structural properties of optimal decision rules are available (see, e.g., [24]), not much is known about the optimal performance. Networks with feedback face the same difficulty. In this paper we consider sensor network architectures that involve feedback: some or all of the sensors have access to compressed summaries of other sensors’ observations. We are interested in characterizing the performance under different architectures, and, in particular, to determine whether the presence of feedback can substantially enhance performance. In the context of a wireless network, feedback can result in unnecessary communication

and computation costs, therefore it is important to quantify the performance gain, if any, that feedback can provide.

A variety of feedback architectures, under a Bayesian formulation, have been studied in [25], [26]. These references show that it is person-by-person optimal for every sensor to use a likelihood ratio quantizer, with thresholds that depend on the feedback messages. However, because of the difficulty of optimizing these thresholds when the number of sensors becomes large, it is difficult to analytically compare the performance of networks with and without feedback. Numerical examples in [26] show that a system with feedback has lower probability of error, as expected. To better understand the asymptotics of the error probability, [27] studies the error probability decay rate under a Neyman-Pearson formulation for two different feedback architectures. For either case, it shows that if the fusion center also has access to the feedback messages, then feedback does not improve the optimal error exponent. References [28], [29] consider the Neyman-Pearson problem in a “daisy-chain” architecture, and obtain a similar result.

In this paper, we revisit the two-message architecture studied in [27], and extend the available results. We also study certain feedback architectures that have not been studied before. Our main contributions are as follows.

- 1) We consider the **two-message full feedback** architecture studied in [27]. Here, each sensor gets to transmit two messages, and the second message can take into account the first messages of all sensors. We resolve an open problem for the Bayesian formulation, by showing that there is no performance gain over the non-feedback case. We also provide a variant of the result of [27] for the Neyman-Pearson case. Our model is different because unlike [27], because we do not restrict the feedback message alphabet to grow at most subexponentially with the number of sensors.
- 2) We also study a new **two-message sequential feedback** architecture. Sensors are indexed, and the second message of a sensor can take into account the first message of all sensors with lower indices. We show that under either the Neyman-Pearson or Bayesian formulation, feedback does not improve the error exponent.

The remainder of this paper is organized as follows. In Section II we define the model, formulate the problems that we will be studying, and provide some background material. In Section III, we study two-message feedback architectures (sequential and full feedback). We offer concluding remarks and discuss open problems in Section IV.

This research was supported, in part, by the Nanyang Technological University Startup Grant, and by the NSF under grant ECCS-0701623. W. P. Tay is with the Nanyang Technological University, Singapore. J. N. Tsitsiklis is with the Laboratory for Information and Decision Systems, MIT, USA. E-mail: wptay@ntu.edu.sg, jnt@mit.edu

II. PROBLEM FORMULATION

We consider a decentralized binary detection problem involving n sensors and a fusion center. Each sensor k observes a random variable X_k taking values in some measurable space $(\mathcal{X}, \mathcal{F})$, and is distributed according to a measure \mathbb{P}_j under hypothesis H_j , for $j = 0, 1$. Under either hypothesis H_j , $j = 0, 1$, the random variables X_k are assumed to be i.i.d. We use \mathbb{E}_j to denote the expectation operator with respect to \mathbb{P}_j .

A. Feedback Architectures

In the two-message feedback architecture (see Figure 1), each sensor k sends a message $Y_k = \gamma_k(X_k)$, which is a quantized version of its observation X_k , to the fusion center. The quantization function is of the form $\gamma_k : \mathcal{X} \mapsto \mathcal{T}$, where \mathcal{T} is the transmission signal set. In most engineering scenarios, \mathcal{T} is assumed to be a finite alphabet, although we do not require this restriction. We denote by Γ the set of allowable quantization functions.

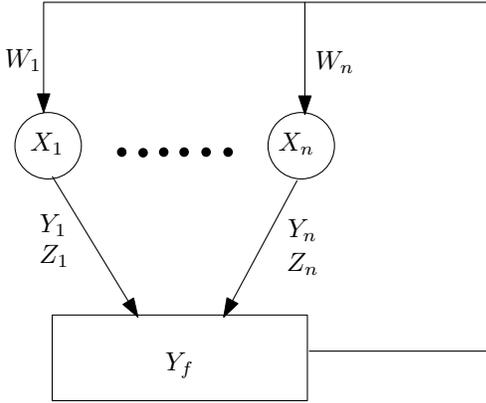


Fig. 1. A two-message architecture.

We assume that the sensors are indexed in order of when they send their messages to the fusion center. We will consider two forms of feedback under the two-message architecture. In the first form, the information $W_k = (Y_1, \dots, Y_{k-1})$ received from sensors $1, \dots, k-1$, is fed back by the fusion center to sensor k . We call this sequential feedback. In the second form of feedback, the fusion center broadcasts the messages (Y_1, \dots, Y_n) to all the sensors. In this case, the additional information received at sensor k is $W_k = (Y_1, \dots, Y_{k-1}, Y_{k+1}, \dots, Y_n)$. We call this full feedback.

In both feedback scenarios, each sensor forms a new second message $Z_k = \delta_k(X_k, W_k)$ based on the additional information W_k , and sends it to the fusion center. Finally, the fusion center makes a decision $Y_f = \gamma_f(Y_1, \dots, Y_n, Z_1, \dots, Z_n)$. The collection $(\gamma_f, \gamma_1, \dots, \gamma_n, \delta_1, \dots, \delta_n)$ is called a strategy. A sequence of strategies, one for each value of n , is called a strategy sequence. Typically, we wish to design strategy sequences that optimize a given objective function.

B. Preliminaries

In this section, we list the basic assumptions that we will be making throughout this paper, and some preliminary

mathematical results that we will apply in our subsequent proofs.

Consider the Radon-Nikodym derivative ℓ_{ij}^X of the measure \mathbb{P}_i with respect to (w.r.t.) the measure \mathbb{P}_j . Informally, this is the likelihood ratio associated with an observation of X . It is a random variable whose value is determined by X ; accordingly, its value should be denoted by a notation such as $\ell_{ij}^X(X)$. However, in order to avoid cluttered expressions, we will abuse notation and just write $\ell_{ij}(X)$. Similarly, we use $\ell_{ij}(X|Y)$ to denote the Radon-Nikodym derivative of the conditional distribution of X given Y under \mathbb{P}_i w.r.t. that under \mathbb{P}_j .

For technical reasons (see [30]), we make the following assumptions.

Assumption 1: The measures \mathbb{P}_0 and \mathbb{P}_1 are absolutely continuous w.r.t. each other. Furthermore, there exists some $\gamma \in \Gamma$ such that $-\mathbb{E}_0[\log \ell_{01}(\gamma(X_1))] < 0 < \mathbb{E}_1[\log \ell_{10}(\gamma(X_1))]$.

Assumption 2: We have $\mathbb{E}_0[\log^2 \ell_{01}(X_1)] < \infty$ and $\mathbb{E}_1[\log^2 \ell_{10}(X_1)] < \infty$.

We will require the following lower bound for the maximum of the Type I and II error probabilities. This bound was first proved in [31] for the case of discrete observation space. The following proposition generalizes the result to a general observation space. The proof is almost identical to that in [31], and is omitted here.¹

Proposition 1: Consider a hypothesis testing problem based on a single observation X with distribution \mathbb{P}_j under hypothesis H_j , $j = 0, 1$. Let $P_{e,j}$ be the probability of error when H_j is true. Let $Z = \log \frac{d\mathbb{P}_1}{d\mathbb{P}_0}(X)$ be the log Radon-Nikodym derivative. For any $s \in \mathbb{R}$, let $\Lambda(s) = \log \mathbb{E}_0[\exp(sZ)]$ be the log-moment generating function of Z . Then, for $s^* \in [0, 1]$ such that $\Lambda'(s^*) = 0$, we have

$$\max(P_{e,0}, P_{e,1}) \geq \frac{1}{4} \exp\left(\Lambda(s^*) - \sqrt{2\Lambda''(s^*)}\right).$$

III. ASYMPTOTIC PERFORMANCE

In this section, we study the Neyman-Pearson and Bayesian formulations of the decentralized detection problem in the two-message architecture. For both sequential and full feedback, and $i, j \in \{0, 1\}$, let the log likelihood ratio at the fusion center be

$$\begin{aligned} \mathcal{L}_{ij}^{(n)} &= \log \ell_{ij}(Y_1, \dots, Y_n, Z_1, \dots, Z_n) \\ &= \sum_{k=1}^n \log \ell_{ij}(Y_k) + \log \ell_{ij}(Z_1, \dots, Z_n | Y_1, \dots, Y_n) \\ &= \sum_{k=1}^n \log \ell_{ij}(Y_k) + \sum_{k=1}^n \log \ell_{ij}(Z_k | Y_k, W_k) \end{aligned}$$

To simplify notation, we let

$$\begin{aligned} \mathcal{L}_{ij}^k(w) &= \log \ell_{ij}(Y_k) + \log \ell_{ij}(Z_k | Y_k, w) \\ &= \log \ell_{ij}(\gamma_k(X_k), \delta_k^w(X_k)), \end{aligned}$$

where $\delta_k^w : \mathcal{X} \mapsto \mathcal{T}$ is a function that depends on the value w .

¹Throughout this paper, we use $f'(s)$ to denote the derivative of f w.r.t. s .

A. Neyman-Pearson Formulation

Let $\alpha \in (0, 1)$ be a given constant. A strategy is admissible if its Type I error probability satisfies $\mathbb{P}_0(Y_f = 1) < \alpha$. Let $\beta_n^* = \inf \mathbb{P}_1(Y_f = 0)$, where the infimum is taken over all admissible strategies. Our objective is to characterize the optimal error exponent $\limsup_{n \rightarrow \infty} (1/n) \log \beta_n^*$, under different feedback architectures.

Let g_p^* be the optimal error exponent of the parallel configuration in which there is no feedback from the fusion center, i.e., each sensor k sends a message $(\gamma_k(X_k), \delta_k(X_k))$ to the fusion center. From [30], the optimal error exponent is

$$g_p^* = - \sup_{(\gamma, \delta) \in \Gamma^2} \mathbb{E}_0 [\log \ell_{01}(\gamma(X_1), \delta(X_1))].$$

Let g_{sf}^* and g_f^* be the optimal error exponents for the sequential feedback and full feedback architectures respectively. Since the sensors can ignore some or all of the feedback messages from the fusion center, we have

$$g_f^* \leq g_{sf}^* \leq g_p^*. \quad (1)$$

We will show that under appropriate but mild assumptions, the inequalities in (1) are equalities. Hence, from an asymptotic view point, both sequential and full feedback results in no gain in detection performance. This is in line with the results in [29], which shows that feedback does not improve the optimal error exponent in a daisy chain architecture. We first show an instructive result that will guide us in the main proofs.

Lemma 1: Suppose that β_n is the Type II error probability of a given strategy. If $R > 0$ and

$$\limsup_{n \rightarrow \infty} \mathbb{P}_0(\mathcal{L}_{01}^{(n)} > nR) < 1 - \alpha,$$

then

$$\liminf_{n \rightarrow \infty} \frac{1}{n} \log \beta_n \geq -R.$$

Proof: We have

$$\begin{aligned} \beta_n &= \mathbb{P}_1(Y_f = 0) \\ &= \mathbb{E}_0 \left[\exp(-\mathcal{L}_{01}^{(n)}) \mathbf{1}_{\{Y_f=0\}} \right] \\ &\geq \mathbb{E}_0 \left[\exp(-\mathcal{L}_{01}^{(n)}) \mathbf{1}_{\{Y_f=0, \mathcal{L}_{01}^{(n)} \leq nR\}} \right] \\ &\geq e^{-nR} \mathbb{P}_0(Y_f = 0, \mathcal{L}_{01}^{(n)} \leq nR). \end{aligned}$$

Therefore,

$$\mathbb{P}_0(Y_f = 0, \mathcal{L}_{01}^{(n)} \leq nR) \leq \beta_n e^{nR}.$$

This upper bound yields

$$1 - \alpha \leq \mathbb{P}_0(Y_f = 0) \leq \beta_n e^{nR} + \mathbb{P}_0(\mathcal{L}_{01}^{(n)} > nR)$$

and

$$\frac{1}{n} \log \beta_n + R \geq \frac{1}{n} \log(1 - \alpha - \mathbb{P}_0(\mathcal{L}_{01}^{(n)} > nR)).$$

The lemma then follows by letting $n \rightarrow \infty$. \blacksquare

Theorem 1: Suppose Assumptions 1-2 hold. Then, the optimal error exponent for the sequential feedback architecture is $g_{sf}^* = g_p^*$. Moreover, there is no loss in optimality if

sensors ignore the feedback messages from the fusion center, and all sensors are constrained to using the same quantization function.

Proof: From (1), we have $g_{sf}^* \leq g_p^*$. To show the reverse inequality, we will first upper bound $\mathbb{E}_0[\mathcal{L}_{01}^k(W_k) | W_k]$ by $R = -g_p^*$, and then apply Lemma 1 to obtain a lower bound for g_{sf}^* . We have for all w ,

$$\begin{aligned} &\mathbb{E}_0 [\mathcal{L}_{01}^k(W_k) | W_k = w] \\ &= \mathbb{E}_0 [\log \ell_{01}(\gamma_k(X_k), \delta_k^w(X_k)) | W_k = w] \\ &\leq \sup_{(\gamma, \delta) \in \Gamma^2} \mathbb{E}_0 [\log \ell_{01}(\gamma(X_1), \delta(X_1))] \\ &= R. \end{aligned} \quad (2)$$

From Assumption 2, there exists some constant $a > 0$ such that

$$\begin{aligned} \text{var}_0(\mathcal{L}_{01}^k(W_k)) &\leq \mathbb{E}_0 \left[\mathbb{E}_0 \left[(\mathcal{L}_{01}^k(W_k))^2 | W_k \right] \right] \\ &\leq a. \end{aligned} \quad (3)$$

Recall that $W_k = (Y_1, \dots, Y_{k-1})$, so we have, for $m < k$,

$$\begin{aligned} &\mathbb{E}_0 [(\mathcal{L}_{01}^m(W_m) - \mathbb{E}_0[\mathcal{L}_{01}^m(W_m) | W_m]) \\ &\quad \cdot (\mathcal{L}_{01}^k(W_k) - \mathbb{E}_0[\mathcal{L}_{01}^k(W_k) | W_k])] \\ &= \mathbb{E}_0 [(\mathcal{L}_{01}^m(W_m) - \mathbb{E}_0[\mathcal{L}_{01}^m(W_m) | W_m]) \\ &\quad \cdot \mathbb{E}_0 [\mathcal{L}_{01}^k(W_k) - \mathbb{E}_0[\mathcal{L}_{01}^k(W_k) | W_k] | W_m]] \\ &= 0. \end{aligned} \quad (4)$$

Let $\epsilon > 0$. From Chebychev's inequality, together with (2), (3), and (4), we obtain

$$\begin{aligned} &\mathbb{P}_0 \left(\mathcal{L}_{01}^{(n)} > n(1 + \epsilon)R \right) \\ &\leq \mathbb{P}_0 \left(\sum_{k=1}^n (\mathcal{L}_{01}^{(n)} - \mathbb{E}_0[\mathcal{L}_{01}^k(W_k) | W_k]) > n\epsilon R \right) \\ &\leq \frac{a}{n\epsilon^2 R^2}. \end{aligned}$$

Letting $n \rightarrow \infty$, we get

$$\lim_{n \rightarrow \infty} \mathbb{P}_0 \left(\mathcal{L}_{01}^{(n)} > n(1 + \epsilon)R \right) = 0.$$

Therefore, applying Lemma 1, we have $g_{sf}^* \geq (1 + \epsilon)g_p^*$. Since ϵ was chosen arbitrarily, we have $g_{sf}^* \geq g_p^*$, and the proof is now complete. \blacksquare

Next, we consider the full feedback architecture. The same architecture has been studied in [27], using the method of types. In the following, we show the same result as [27], i.e., there is no gain from the feedback messages asymptotically, but without the constraint that the feedback messages have an alphabet that grows at most subexponentially fast.

For technical reasons, we will require the following additional assumption.

Assumption 3: Let $b(\lambda) = \log \mathbb{E}_0[\exp(\lambda \log \ell_{01}(X_1))]$. There exists $s > 0$ such that $b(s) < \infty$.

Note that since $b(\cdot)$ is non-decreasing on $[0, s]$ (see Lemma 2.2.5 of [32]), and therefore Assumption 3 implies that $b(\lambda) < \infty$ for all $\lambda \in [0, s]$.

Suppose that a strategy sequence has been fixed. Let $\varphi_n(\lambda) = \log \mathbb{E}_0[\exp(\lambda \mathcal{L}_{01}^{(n)}/n)]$. The Fenchel-Legendre transform of φ_n is $\Phi_n(t) = \sup_{\lambda \in \mathbb{R}} \{\lambda t - \varphi_n(\lambda)\}$. We first show some properties of φ_n , and use these properties together with Lemma 1 to prove that similar to the sequential feedback architecture, feedback does not improve the asymptotic performance of a full feedback architecture. Denote by $\varphi'_n(\lambda)$ the derivative of φ_n w.r.t. to λ .

Lemma 2: Suppose Assumption 3 holds, and let s be as in the assumption. Suppose $\lambda \in [0, s)$. Then, for $n > 1$, we have $0 \leq \varphi_n(\lambda) \leq nb(\lambda/n)$, $0 \leq \varphi'_n(\lambda) \leq C_1$, and $0 \leq \varphi''_n(\lambda) \leq C_2$ for some constants $C_1, C_2 > 0$, independent of the strategy used and λ . Furthermore, $\varphi'_n(0) \leq -g_p^*$ for all n .

Proof: Since φ_n is a convex function [32] with $\varphi_n(-1) = \varphi_n(0) = 0$, $\varphi_n(\lambda) \geq 0$ for $\lambda \geq 0$. Using Jensen's inequality, we obtain

$$\begin{aligned} \frac{1}{n} \varphi_n(\lambda) &= \frac{1}{n} \log \mathbb{E}_0 \left[\exp(\lambda \mathcal{L}_{01}^{(n)}/n) \right] \\ &\leq \frac{1}{n} \log \mathbb{E}_0 \left[\exp((\lambda/n) \log \ell_{01}(X_1, \dots, X_n)) \right] \\ &= \frac{1}{n} \sum_{k=1}^n \log \mathbb{E}_0 \left[\exp((\lambda/n) \log \ell_{01}(X_k)) \right] \\ &= b(\lambda/n) < \infty. \end{aligned} \quad (5)$$

From (5), it can be shown that φ_n is twice differentiable throughout $[0, s)$ (see e.g. Example A.5.2 of [33]). To simplify the notation in the following proof, let $\ell_{01}^{(n)} = \ell_{01}(Y_1, \dots, Y_n, Z_1, \dots, Z_n)$ and $\ell_{01,n} = \ell_{01}(X_1, \dots, X_n)$. Since $\varphi_n(\lambda) \geq 0$, we have

$$\begin{aligned} \varphi'_n(\lambda) &= \frac{1}{n} e^{-\varphi_n(\lambda)} \mathbb{E}_0 \left[e^{(\lambda/n) \mathcal{L}_{01}^{(n)}} \mathcal{L}_{01}^{(n)} \right] \\ &\leq \frac{1}{n} \mathbb{E}_0 \left[(\ell_{01}^{(n)})^{\lambda/n} \log \ell_{01}^{(n)} \right] \\ &= \frac{1}{n} \mathbb{E}_1 \left[(\ell_{01}^{(n)})^{\lambda/n+1} \log \ell_{01}^{(n)} \right] \\ &\leq \frac{1}{n} \mathbb{E}_1 \left[\phi(\ell_{01}^{(n)}) \right], \end{aligned} \quad (6)$$

where $\phi(x) = (x^{\lambda/n+1} \log x) \mathbf{1}_{\{x \geq 1\}}$ is a convex function. Since $\ell_{01}^{(n)} = \mathbb{E}_1[\ell_{01,n} | \mathcal{F}]$, where \mathcal{F} is the σ -field generated by Y_k and Z_k , $k = 1, \dots, n$, we get from Jensen's inequality that the R.H.S. of (6) is bounded above by

$$\begin{aligned} &\frac{1}{n} \mathbb{E}_1 \left[\phi(\ell_{01,n}) \right] \\ &\leq \frac{1}{n} \mathbb{E}_0 \left[(\ell_{01,n})^{\lambda/n} \log \ell_{01,n} \right] \\ &\leq \frac{1}{n} \mathbb{E}_0 \left[(\ell_{01,n})^{\lambda/n} \sum_{k=1}^n |\log \ell_{01}(X_k)| \right] \\ &= \mathbb{E}_0 \left[(\ell_{01}(X_1))^{\lambda/n} \log \ell_{01}(X_1) \right] \mathbb{E}_0 \left[(\ell_{01}(X_1))^{\lambda/n} \right]^{n-1} \\ &\leq \mathbb{E}_0 \left[(\ell_{01}(X_1))^{2\lambda/n} \right]^{1/2} \mathbb{E}_0 \left[\log^2 \ell_{01}(X_1) \right]^{1/2} \\ &\quad \cdot \mathbb{E}_0 \left[(\ell_{01}(X_1))^{(1-1/n)\lambda} \right] \end{aligned} \quad (7)$$

where the last equality is because the X_i are i.i.d., and the last inequality follows from the Cauchy-Schwarz inequality and Jensen's inequality. From Assumption 3, the R.H.S. of (7) is

bounded by some constant C_1 , independent of n and λ . The proof for the bound on the second derivative is similar, and is omitted. Finally, the bound for $\varphi'_n(0) = (1/n) \mathbb{E}_0 \left[\mathcal{L}_{01}^{(n)} \right]$ follows from the same proof as for (2). The proof of the lemma is now complete. ■

Lemma 3: Suppose Assumptions 1-3 hold, and $t > r = \limsup_{n \rightarrow \infty} \varphi'_n(0)$. Then, there exist N_0 and $\eta > 0$ such that for all $n \geq N_0$, $\Phi_n(t) \geq \eta$.

Proof: Let $\epsilon = t - r > 0$, and let $\lambda = \min\{s/2, \epsilon/(2C_2)\} > 0$. Then, from Lemma 2, we have $\varphi_n(\lambda) \leq (\varphi'_n(0) + \lambda C_2)\lambda$. There exists N_0 such that for all $n \geq N_0$, $\varphi'_n(0) \leq r + \epsilon/4$. Therefore, we get

$$\begin{aligned} \varphi_n(\lambda) &\leq (r + 3\epsilon/4)\lambda \\ &= (t - \epsilon/4)\lambda. \end{aligned}$$

Therefore, $\Phi_n(t) \geq \lambda t - \varphi_n(\lambda) \geq \epsilon\lambda/4$. The proof of the lemma is now complete. ■

Finally, we can prove the following result.

Theorem 2: Suppose Assumptions 1-3 hold. Then, in the full feedback architecture, there is no loss in optimality if sensors ignore the feedback messages from the fusion center, i.e., $g_f^* = g_p^*$. Moreover, there is no loss in optimality if all sensors are constrained to using the same quantization function.

Proof: Let $\epsilon > 0$ and $t = -(1 + \epsilon)g_p^*$. From Lemma 2, $\limsup_{n \rightarrow \infty} \varphi'_n(0) \leq -g_p^*$. Applying the Chernoff bound and Lemma 3, we have for n sufficiently large,

$$\frac{1}{n} \log \mathbb{P}_0 \left(\mathcal{L}_{01}^{(n)} > nt \right) \leq -\Phi_n(t) \leq -\eta,$$

where $\eta > 0$ is as chosen in Lemma 3. Therefore we can apply Lemma 1 to get $g_f^* \geq (1 + \epsilon)g_p^*$. Taking $\epsilon \rightarrow 0$ and applying (1), we obtain the theorem. ■

B. Bayesian Formulation

Let the prior probability of hypothesis H_j be $\pi_j > 0$, $j = 0, 1$. Given a strategy, the probability of error at the fusion center is $P_e = \pi_0 \mathbb{P}_0(Y_f = 1) + \pi_1 \mathbb{P}_1(Y_f = 0)$. Let P_e^* be the minimum probability of error, over all strategies. We seek to characterize the optimal error exponent

$$\limsup_{n \rightarrow \infty} \frac{1}{n} \log P_e^*.$$

From [30], the optimal error exponent for the parallel configuration without any feedback is given by

$$\mathcal{E}_p^* = \inf_{(\gamma, \delta) \in \Gamma} \min_{\lambda \in [0, 1]} \log \mathbb{E}_0 \left[\exp(\lambda \log \ell_{10}(\gamma(X_1), \delta(X_1))) \right].$$

Similar to the Neyman-Pearson case, we let \mathcal{E}_{sf}^* and \mathcal{E}_f^* denote the optimal error exponent for the sequential feedback and full feedback architectures respectively. We first study the full feedback architecture, and infer results about the sequential feedback architecture.

Let $\psi_n(s) = \log \mathbb{E}_0 \left[\exp(s \mathcal{L}_{10}^{(n)}) \right]$ and $\Psi_n(t) = \sup_{s \in \mathbb{R}} \{st - \psi_n(s)\}$. We will require the following results in our proofs.

Lemma 4: Suppose Assumptions 1 and 2 hold.

- (i) For all $s \in [0, 1]$, we have $\mathbb{E}_0[\log \ell_{10}(X_1)] \leq \psi'_n(s)/n \leq \mathbb{E}_1[\log \ell_{10}(X_1)]$.
- (ii) Let t be such that for all n , there exists $s_n \in (0, 1)$ with $\psi'_n(s_n) = t$. Then, there exists a constant C such that for all n , we have $\psi''_n(s_n) \leq nC$.
- (iii) For all $s \in [0, 1]$, we have $\psi_n(s) \geq n\mathcal{E}_p^*$.

Proof: (Outline) Claims (i) and (ii) follow from calculus. We omit their proofs here. In the following, we show claim (iii). Let $\bar{\lambda} = \lambda/n$ and $Y_1^n = (Y_1, \dots, Y_n)$. We have

$$\begin{aligned} \psi_n(\lambda) &= \log \mathbb{E}_0 [(\ell_{10}(Y_1^n))^{\bar{\lambda}} \\ &\quad \cdot \mathbb{E}_0 \left[\prod_{k=1}^n (\ell_{10}(Z_k|Y_k, W_k))^{\bar{\lambda}} \mid Y_1^n \right]]. \end{aligned} \quad (8)$$

Recall that $Z_k = \delta_k^{W_k}(X_k)$, where $W_k = (Y_1^{k-1}, Y_{k+1}^n)$ in the full feedback configuration. Consider the inner expectation on the R.H.S. of (8). For $\epsilon > 0$, we have

$$\begin{aligned} &\mathbb{E}_0 \left[\prod_{k=1}^n (\ell_{10}(Z_k|Y_k, W_k))^{\bar{\lambda}} \mid Y_1^n \right] \\ &= \mathbb{E}_0 \left[\prod_{k=1}^n (\ell_{10}(\delta_k^{W_k}(X_k)|Y_k))^{\bar{\lambda}} \mid Y_1^n \right] \\ &= \prod_{k=1}^n \mathbb{E}_0 \left[(\ell_{10}(\delta_k^{W_k}(X_k)|Y_k))^{\bar{\lambda}} \mid Y_1^n \right] \\ &\geq \prod_{k=1}^n \left(\mathbb{E}_0 \left[(\ell_{10}(\delta_k^{Y_k}(X_k)|Y_k))^{\bar{\lambda}} \mid Y_k \right] - \epsilon \right), \end{aligned}$$

where $\delta_k^{Y_k} \in \Gamma$ is a function depending on the value of Y_k , and is such that

$$\begin{aligned} &\mathbb{E}_0 \left[(\ell_{10}(\delta_k^{Y_k}(X_k)|Y_k))^{\bar{\lambda}} \mid Y_k \right] \\ &\leq \inf_{\delta \in \Gamma} \mathbb{E}_0 \left[(\ell_{10}(\delta(X_k)|Y_k))^{\bar{\lambda}} \mid Y_k \right] + \epsilon. \end{aligned}$$

From (8), we obtain

$$\begin{aligned} \psi_n(\lambda) &\geq \log \mathbb{E}_0 [(\ell_{10}(Y_1^n))^{\bar{\lambda}} \\ &\quad \cdot \prod_{k=1}^n \left(\mathbb{E}_0 \left[(\ell_{10}(\delta_k^{Y_k}(X_k)|Y_k))^{\bar{\lambda}} \mid Y_k \right] - \epsilon \right)] \\ &= \sum_{k=1}^n \log \mathbb{E}_0 [(\ell_{10}(Y_k))^{\bar{\lambda}} \\ &\quad \cdot \left(\mathbb{E}_0 \left[(\ell_{10}(\delta_k^{Y_k}(X_k)|Y_k))^{\bar{\lambda}} \mid Y_k \right] - \epsilon \right)] \\ &\geq \sum_{k=1}^n \log \left(\mathbb{E}_0 \left[(\ell_{10}(Y_k, \delta_k^{Y_k}(X_k)))^{\bar{\lambda}} \right] - \epsilon \right), \end{aligned} \quad (9)$$

where we used the inequality $\mathbb{E}_0 [(\ell_{10}(Y_k))^{\bar{\lambda}}] \leq 1$ in (9). Recall that $Y_k = \gamma_k(X_k)$. We can define $\xi_k \in \Gamma^2$ such that $\xi_k(X_k) = (\gamma_k(X_k), \delta_k(X_k))$, where $\delta_k(X_k) = \delta_k^u(X_k)$ iff $\gamma_k(X_k) = u \in \mathcal{T}$. From (9), we obtain the bound

$$\begin{aligned} \psi_n(\lambda) &\geq \sum_{k=1}^n \log \left(\mathbb{E}_0 \left[(\ell_{10}(\xi_k(X_k)))^{\bar{\lambda}} \right] - \epsilon \right) \\ &\geq n \log \left(\inf_{\xi \in \Gamma^2} \mathbb{E}_0 \left[(\ell_{10}(\xi(X_1)))^{\bar{\lambda}} \right] - \epsilon \right) \end{aligned}$$

Since ϵ is arbitrary, we have

$$\begin{aligned} \psi_n(\lambda) &\geq n \inf_{\xi \in \Gamma^2} \log \mathbb{E}_0 \left[(\ell_{10}(\xi(X_1)))^{\bar{\lambda}} \right] \\ &\geq n\mathcal{E}_p^*, \end{aligned}$$

and the lemma is proved. \blacksquare

Theorem 3: Suppose Assumptions 1 and 2 hold. Then $\mathcal{E}_f^* = \mathcal{E}_p^*$. Moreover, there is no loss in optimality if sensors are constrained to using the same quantization function, which ignore the feedback messages from the fusion center.

Proof: It is clear that $\mathcal{E}_f^* \leq \mathcal{E}_p^*$. To show the reverse bound, we make use of Proposition 1. Let the conditional probability of error under H_j be $P_{n,j}$ for $j = 0, 1$. Let $s_n^* = \arg \min_{s \in (0,1)} \psi_n(s)$ so that $\psi'_n(s_n^*) = 0$. From Proposition 1, we have

$$\begin{aligned} \max_{j=0,1} P_{n,j} &\geq \frac{1}{4} \exp \left(\psi_n(s_n^*) - \sqrt{2\psi''_n(s_n^*)} \right) \\ &\geq \exp(\psi_n(s_n^*) - C\sqrt{n}) \\ &\geq \exp(n\mathcal{E}_p^* - C\sqrt{n}) \end{aligned}$$

where C is some constant. In the above, the penultimate inequality follows from Lemma 4(ii), and the last inequality from Lemma 4(iii). Letting $n \rightarrow \infty$, we have

$$\begin{aligned} \limsup_{n \rightarrow \infty} \frac{1}{n} \log P_e &= \limsup_{n \rightarrow \infty} \frac{1}{n} \log \max_{j=0,1} P_{n,j} \\ &\geq \mathcal{E}_p^*. \end{aligned}$$

This implies that $\mathcal{E}_f^* \geq \mathcal{E}_p^*$, and the theorem is proven. \blacksquare

Since the sequential feedback configuration can perform no better than the full feedback architecture, and no worse than the parallel configuration, we have the following result.

Theorem 4: Suppose Assumptions 1 and 2 hold. Then $\mathcal{E}_{sf}^* = \mathcal{E}_p^*$. Moreover, there is no loss in optimality if sensors are constrained to using the same quantization function, which ignore the feedback messages from the fusion center.

IV. CONCLUSION

We have studied the two-message feedback architecture, in which each sensor has access to compressed summaries of some or all other sensors' first messages to the fusion center. In the sequential feedback architecture, each sensor has access to the first messages of those sensors that communicate with the fusion center before it. In the full feedback architecture, each sensor has access to the first messages of every other sensor. In both architectures, and under both the Neyman-Pearson and Bayesian formulations, we show that the optimal error exponent is not improved by the feedback messages. Our results suggest that in the regime of a large number of sensors, the performance gain due to feedback does not justify the increase in communication and computation costs incurred in a feedback architecture.

Future research directions include studying the effect of feedback in architectures where all sensors send only one message to the fusion center, and where a group of sensors has access to the messages of the other sensors. Other more complex feedback architectures like hierarchical networks with

feedback messages from one level to the next are also important future research directions. We would also like to consider the impact of feedback on distributed multiple hypothesis testing and parameter estimation.

REFERENCES

- [1] R. R. Tenney and N. R. Sandell, "Detection with distributed sensors," *IEEE Trans. Aerosp. Electron. Syst.*, vol. 17, no. 4, pp. 501–510, 1981.
- [2] Z. Chair and P. K. Varshney, "Optimal data fusion in multiple sensor detection systems," *IEEE Trans. Aerosp. Electron. Syst.*, vol. 22, no. 1, pp. 98–101, 1986.
- [3] G. Polychronopoulos and J. N. Tsitsiklis, "Explicit solutions for some simple decentralized detection problems," *IEEE Trans. Aerosp. Electron. Syst.*, vol. 26, no. 2, pp. 282–292, 1990.
- [4] J. N. Tsitsiklis, "Decentralized detection," *Advances in Statistical Signal Processing*, vol. 2, pp. 297–344, 1993.
- [5] W. W. Irving and J. N. Tsitsiklis, "Some properties of optimal thresholds in decentralized detection," *IEEE Trans. Autom. Control*, vol. 39, no. 4, pp. 835–838, 1994.
- [6] R. Viswanathan and P. K. Varshney, "Distributed detection with multiple sensors: part I - fundamentals," *Proc. IEEE*, vol. 85, no. 1, pp. 54–63, 1997.
- [7] B. Chen and P. K. Varshney, "A Bayesian sampling approach to decision fusion using hierarchical models," *IEEE Trans. Signal Process.*, vol. 50, no. 8, pp. 1809–1818, Aug. 2002.
- [8] B. Chen and P. K. Willett, "On the optimality of the likelihood-ratio test for local sensor decision rules in the presence of nonideal channels," *IEEE Trans. Inf. Theory*, vol. 51, no. 2, pp. 693–699, Feb. 2005.
- [9] A. Kashyap, "Comments on 'On the optimality of the likelihood-ratio test for local sensor decision rules in the presence of nonideal channels,'" *IEEE Trans. Inf. Theory*, vol. 52, no. 3, pp. 1274–1275, Mar. 2006.
- [10] B. Liu and B. Chen, "Channel-optimized quantizers for decentralized detection in sensor networks," *IEEE Trans. Inf. Theory*, vol. 52, no. 7, pp. 3349–3358, Jul. 2006.
- [11] R. Viswanathan, S. C. A. Thomopoulos, and R. Tumuluri, "Optimal serial distributed decision fusion," *IEEE Trans. Aerosp. Electron. Syst.*, vol. 24, no. 4, pp. 366–376, 1988.
- [12] Z. B. Tang, K. R. Pattipati, and D. L. Kleinman, "Optimization of detection networks: part I - tandem structures," *IEEE Trans. Syst., Man, Cybern.*, vol. 21, no. 5, pp. 1044–1059, 1991.
- [13] J. D. Papastavrou and M. Athans, "Distributed detection by a large team of sensors in tandem," *IEEE Trans. Aerosp. Electron. Syst.*, vol. 28, no. 3, pp. 639–653, 1992.
- [14] W. P. Tay, J. N. Tsitsiklis, and M. Z. Win, "On the sub-exponential decay of detection error probabilities in long tandems," *IEEE Trans. Inf. Theory*, vol. 54, no. 10, pp. 4767–4771, Oct. 2008.
- [15] L. K. Ekchian and R. R. Tenney, "Detection networks," vol. 21, Dec. 1982, pp. 686–691.
- [16] A. R. Reibman and L. W. Nolte, "Design and performance comparison of distributed detection networks," *IEEE Trans. Aerosp. Electron. Syst.*, vol. 23, no. 6, pp. 789–797, 1987.
- [17] Z. B. Tang, K. R. Pattipati, and D. L. Kleinman, "Optimization of detection networks: part II - tree structures," *IEEE Trans. Syst., Man, Cybern.*, vol. 23, no. 1, pp. 211–221, 1993.
- [18] A. Pete, K. Pattipati, and D. Kleinman, "Optimization of detection networks with multiple event structures," *IEEE Trans. Autom. Control*, vol. 39, no. 8, pp. 1702–1707, 1994.
- [19] Y. Lin, B. Chen, and P. K. Varshney, "Decision fusion rules in multi-hop wireless sensor networks," *IEEE Trans. Aerosp. Electron. Syst.*, vol. 41, no. 2, pp. 475–488, Apr. 2005.
- [20] W. P. Tay, J. N. Tsitsiklis, and M. Z. Win, "On the impact of node failures and unreliable communications in dense sensor networks," *IEEE Trans. Signal Process.*, vol. 56, no. 6, pp. 2535–2546, Jun. 2008.
- [21] —, "Data fusion trees for detection: Does architecture matter?" *IEEE Trans. Inf. Theory*, vol. 54, no. 9, pp. 4155–4168, Sep. 2008.
- [22] —, "Bayesian detection in bounded height tree networks," *IEEE Trans. Signal Process.*, vol. 57, no. 10, pp. 4042–4051, Oct. 2009.
- [23] J. N. Tsitsiklis and M. Athans, "On the complexity of decentralized decision making and detection problems," *IEEE Trans. Autom. Control*, vol. 30, pp. 440–446, 1985.
- [24] O. Kreidl and A. Willsky, "An efficient message-passing algorithm for optimizing decentralized detection networks," *IEEE Trans. Autom. Control*, vol. 55, no. 3, pp. 563–578, Mar. 2010.
- [25] S. Alhakeem and P. K. Varshney, "A unified approach to the design of decentralized detection systems," *IEEE Trans. Aerosp. Electron. Syst.*, vol. 31, no. 1, pp. 9–20, 1995.
- [26] S. Alhakeem and P. Varshney, "Decentralized Bayesian detection with feedback," *IEEE Trans. Syst., Man, Cybern.*, vol. 26, no. 4, pp. 503–513, Jul. 1996.
- [27] H. Shalaby and A. Papamarcou, "A note on the asymptotics of distributed detection with feedback," *Information Theory, IEEE Transactions on*, vol. 39, no. 2, pp. 633–640, Mar. 1993.
- [28] S. I. Zoumpoulis, "Decentralized detection in sensor network architectures with feedback," Master's thesis, Massachusetts Institute of Technology, June 2009.
- [29] O. P. Kreidl, J. N. Tsitsiklis, and S. I. Zoumpoulis, "On decentralized detection with partial information sharing among sensors," *IEEE Trans. Signal Process.*, 2011, in press.
- [30] J. N. Tsitsiklis, "Decentralized detection by a large number of sensors," *Math. Control, Signals, Syst.*, vol. 1, pp. 167–182, 1988.
- [31] C. E. Shannon, R. G. Gallager, and E. R. Berlekamp, "Lower bounds to error probability for coding on discrete memoryless channels, I," *Information and Control*, vol. 10, pp. 65–103, 1967.
- [32] A. Dembo and O. Zeitouni, *Large Deviations Techniques and Applications*. New York, NY: Springer-Verlag, 1998.
- [33] R. Durrett, *Probability: Theory and Examples*, 4th ed. New York, NY: Duxbury Press, 2010.