

## Lecture 2 – Median trick, Distinct Count, Impossibility Results

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## 1 Introduction

Today's lecture has three main topics that we'll go through, i.e. Median Trick (from previous lecture), Distinct element count and Impossibility Results.

## 2 Median Trick

So far, we have an algorithm  $A$  which estimates in correct range of  $\epsilon$  with probability  $\geq 0.9$ . Our new algorithm  $A^*$  will output in range of  $\epsilon$  with probability  $1 - \delta$ . Algorithm:

- Repeat  $A$  for  $m = O(\log(1/\delta))$  times
- Take median of all the  $m$  answers.

To prove the correctness, we'll use Chernoff/Hoeffding bounds.

**Definition 1** (Chernoff/Hoeffding Bound). *Let  $X_1, X_2, \dots, X_m$  be independent random variables  $\in \{0, 1\}$ ,  $\mu = E[\sum_i X_i]$ ,  $\epsilon \in [0, 1]$ . Then  $Pr[|\sum_i X_i - \mu| > \epsilon\mu] \leq 2e^{-\epsilon^2\mu/3}$*

Define  $X_i = 1$  iff the  $i^{\text{th}}$  answer of  $A$  is correct (i.e. estimated value of  $A$  lies in correct range).

**Claim 2.**  $E[X_i] = 0.9$ , and  $E[\mu] = 0.9m$

*Proof.* Since  $A$  is correct with probability 0.9,  $E[X_i] = 0.9$ . And  $E[\mu] = 0.9m$  due to linearity of expectation.  $\square$

**Claim 3.** *New algorithm  $A^*$  is correct when  $\sum_i X_i > 0.5m$*

*Proof.* Since we are considering median value to be our answer, if more than half the trials of  $A$  are correct, algorithm  $A^*$  is also correct.  $\square$

**Claim 4.** *To prove,  $Pr[\sum_i X_i \geq 0.5m] \geq 1 - \delta$  or  $Pr[\sum_i X_i < 0.5m] < \delta$*

*Proof.*

$$\begin{aligned}
 Pr[\sum_i X_i < 0.5m] &= Pr[\sum_i X_i - 0.9m < -0.4m] \\
 &\leq Pr[|\sum_i X_i - \mu| > 0.4m] \\
 &= Pr[|\sum X_i - \mu| > 0.4/0.9\mu]
 \end{aligned} \tag{1}$$

Using Chernoff bound,

$$\begin{aligned} &\leq e^{-c*0.9m} \\ &< \delta \end{aligned} \tag{2}$$

Above equation holds for  $m = O(\log(1/\delta))$  □

### 3 Distinct Elements

Given, a stream of size  $m$  containing numbers from  $[n]$ , we have to approximate the number of elements with non-zero frequency. To calculate the exact value the space required:

- $O(n)$  bits. (maintain a vector of length  $n$ ).
- $O(m \log(n))$  bits. (save  $m$  numbers, each taking  $\log(n)$  bits).

Since, this complexity is not feasible as  $m, n$  can be very large, we'll look at algorithm for approximating the distinct count value.

#### 3.0.1 Hash Function

- $h : [n] \rightarrow [0, 1]$
- $h(i)$  is uniformly distributed in  $[0, 1]$ .

#### 3.1 Algorithm [Flajolet-Martin 1985]

We maintain a variable  $z$ .

1. Initialize  $z = 1$ .
2. Whenever  $i$  is encountered:  $z = \min(z, h(i))$
3. When done, output  $1/z - 1$ .

Now, we'll prove the algorithm works in a similar fashion followed in previous lecture. Let  $d$  be number of distinct elements.

**Claim 5.**  $E[z] = d + 1$

*Proof.*  $z$  is the minimum of  $d$  random numbers in  $[0, 1]$ . Pick another random number  $a \in [0, 1]$ . The probability  $a < z$ :

1. exactly  $z$
2. probability it's smallest among  $d + 1$  reals :  $1/(d + 1)$

Equating these two, one can prove the claim. □

**Claim 6.**  $var[z] \leq 2/d^2$

*Proof.* It can be done in a similar fashion described in previous lecture. □

### 3.1.1 $(1 + \epsilon)$ approximation Algorithm

We can take  $Z = (z_1 + z_2 + \dots + z_k)/k$  for independent  $z_1, \dots, z_k$

### 3.2 Alternate Algorithm: Bottom-k

Instead of just use the minimum value of hash function for  $i$  inputs, we'll maintain the  $k$  smallest hashes seen.

1. Initialize  $(z_1, z_2, \dots, z_k) = 1$ .
2. Keep  $k$  smallest hashes seen, s.t.  $z_1 \leq z_2 \leq \dots \leq z_k$
3. When done, output  $\hat{d} = k/z_k$

**Claim 7.** *The following claims are stated:*

- $Pr[\hat{d} > (1 + \epsilon)d] \leq 0.05$
- $Pr[\hat{d} < (1 - \epsilon)d] \leq 0.05$
- Overall probability that  $\hat{d}$  outside range is at most 0.1

*Proof.* To compute  $Pr[\hat{d} > (1 + \epsilon)d]$ :

- Define  $X_i = 1$  iff  $h(i) < \frac{k}{(1 + \epsilon)d}$
- Then  $\hat{d} > (1 + \epsilon)d$  iff  $\sum_i X_i > k$
- if  $\sum_i X_i > k$   
 $\iff \exists$  at least  $k$  numbers for which  $h(i) < \frac{k}{(1 + \epsilon)d}$   
 $\iff z_k < \frac{k}{(1 + \epsilon)d} \iff \frac{k}{z_k} > (1 + \epsilon)d \iff \hat{d} > (1 + \epsilon)d$  (3)

- $E[X_i] = \frac{k}{(1 + \epsilon)d}$   
 $E[\sum_i X_i] = dE[X_i] = \frac{k}{1 + \epsilon}$   
 $\text{var}[\sum_i X_i] = d\text{var}[X_i] \leq dE[X_i^2] \leq \frac{k}{1 + \epsilon} \leq k$   
 (Since  $X_i \in \{0, 1\}$ ,  $E[X_i^2] = E[X_i]$ )
- By Chebyshev:  $Pr[|\sum X_i - \frac{k}{1 + \epsilon}| > \sqrt{20k}] \leq 0.05 \implies Pr[\sum X_i > \frac{k}{1 + \epsilon} + \sqrt{20k}] \leq 0.05$

- (For  $\epsilon < 1/2$  and  $k = c/\epsilon^2$ )  
 $\frac{k}{1 + \epsilon} + \sqrt{20k} \leq k(1 - \epsilon + \epsilon^2) + \sqrt{20k}$  (Taylor Series Expansion)  
 $\leq k - k\epsilon/2 + 5\sqrt{c}/\epsilon = k - c/2\epsilon + 5\sqrt{c}/\epsilon$   
 $< k$  where  $c > 100$

– Since  $k > \frac{k}{1+\epsilon} + \sqrt{20k}$  in our case and  $\Sigma X_i$  is monotonically increasing,  $Pr[\Sigma X_i > k] \leq Pr[\Sigma X_i > \frac{k}{1+\epsilon} + \sqrt{20k}] \leq 0.05$

□

### 3.3 Hash functions in stream

The hash function we used has two practical issues: (1) the return value should be a real number. (2) how do we store it?

Discretization can solve the first issue. Instead of all the real numbers in  $[0, 1]$ , we use hash function with range  $\{0, \frac{1}{M}, \frac{2}{M}, \frac{3}{M}, \dots, 1\}$ . For large  $M \gg n^3$ , the probability that  $d \leq n$  random numbers collide is at most  $\frac{1}{n}$ .

For the second issue, we use pairwise independent function instead of independent function.

**Definition 8.**  $h : [n] \rightarrow \{1, 2, \dots, M\}$  is pairwise independent if for all  $i \neq j$  and  $a, b \in [M]$ ,  $Pr[h(i) = a \wedge h(j) = b] = \frac{1}{M^2}$

It works because in previous calculation, we only care about pairs. We defined  $X_i = 1$  iff  $h(i)$  is small than a threshold, then we computed  $\text{var}[\Sigma X_i] = E[(\Sigma X_i)^2] - E[\Sigma X_i]^2 = E[X_1X_1 + X_1X_2 + \dots] - E[\Sigma X_i]^2$ . Notice that  $E[X_iX_j]$  is the same for fully random  $h$  and pairwise independent  $h$ .

**Example 9** (Construct a pairwise independent hash). Assume  $M$  is a prime number (if not, we can always pick a larger  $M$  that is a prime number). We pick  $p, q \in \{0, 1, 2, \dots, M-1\}$  and the hash function  $h(i) = pi + q \pmod M$ . In this construction we only need  $O(\log M) = O(\log n)$  space (to store  $p, q, M$ ).

*Proof.*  $h(i) = a, h(j) = b$  is equivalent to  $pi + q \equiv a, pj + q \equiv b$ . So  $p(i - j) \equiv a - b$  and  $p \equiv (a - b)(i - j)^{-1}, q \equiv a - pi$ . Since  $M$  is a prime number, the unique inverse implies that there is only one pair  $(p, q)$  satisfies it. And the probability that pair is chosen is exactly  $\frac{1}{M^2}$ . □

## 4 Impossibility Results

We have used both approximation and randomization to solve the distinct counting problem with space much less than  $\min(m, n)$ . Now we are wondering: can we omit either approximation or randomization to achieve the same space efficiency? The answer is no.

### 4.1 Deterministic Exact Won't Work

First, we will show that there is no deterministic (no randomization) and exact (no approximation) way to solve it.

Suppose there do exists a deterministic and exact algorithm  $A$  and an estimator function  $R$  that use space  $s \ll n, m$ . That is, for a given integer stream, we first run the algorithm  $A$  on the stream. As the stream goes  $A$  will return middle memory steps, and we obtain the final memory state  $\sigma$  after the stream ends. Then we apply  $R$  on  $\sigma$  to obtain our estimator  $\hat{d}$ . Since both  $A$  and  $R$  are deterministic and exact,  $\hat{d}$  must equals to the distinct count for the stream.

We now build a binary representation  $x$  of the stream with the following rules: (1)  $x \in \{0, 1\}^n$ , (2)  $i$  in stream iff  $x_i = 1$ . For example, if 1, 3, 5, 6, 7 are in the stream and 2, 4 are not,  $x$  will start with

1, 0, 1, 0, 1, 1, 1. Notice that each stream has a corresponding representation and streams containing different numbers have different representations.

**Claim 10.** *We can recover the  $x$  of the stream given the memory state  $\sigma$*

*Proof.* Denote  $d = R(\sigma)$  be the original estimator. Now we treat  $\sigma$  as a middle snapshot of the memory and add integer  $i$  as the next element of the stream. Now  $A$  will return another memory state  $\sigma'$ , and  $d' = R(\sigma')$  will be our new estimator. If  $d' = d$ ,  $i$  must have appeared in the stream before since  $A$  and  $R$  are deterministic and exact. Similarly, if  $d' > d$ ,  $i$  must have not appeared in the stream before. Using this method with  $i = 1, 2, 3 \dots$  and we can recover the  $x$ .  $\square$

Since we can recover  $x$  from  $\sigma$ , we can treat  $\sigma$  as an encoding of a string  $x$  of length  $n$ . But  $\sigma$  has only  $s \ll n$  bits! Furthermore, we can treat  $A$ , the function that produces  $\sigma$ , as a function with domain  $\{0, 1\}^n$  and  $\{0, 1\}^s$ . We can see that  $A$  must be injective because if  $A(x) = A(x') = \sigma$ , the recoverability implies  $x = x'$ .

Hence  $s \geq n$ . Which implies that there is no deterministic and exact algorithm  $A$  and an estimator function  $R$  that use space  $s \ll n, m$ .

## 4.2 Deterministic Approx. Won't Either

We can use the similar strategy to prove that deterministic approx. won't work. We pick  $T \subset \{0, 1\}^n$  that satisfies the following conditions: (1) for all distinct  $x, y \in T$ , the number of digits  $i$  that  $y_i = 1$  and  $x_i = 0$  should  $\geq \frac{n}{6}$ . (2)  $|T| \geq 2^{\Omega(n)}$ . Now we use algorithm  $A$  to encode an input  $x$  into  $\sigma = A(x)$  and our estimator would be  $\hat{d} = R(\sigma)$ .

Now we want to recover  $x$  based on  $\sigma$ , as what we have done in the last section. For a given  $\sigma$  and any  $y \in T$ , we append  $y$  to the stream and apply  $A$  on it, and  $A$  will return a memory state  $\sigma'$ . Using  $\sigma'$  we have new estimator  $\hat{d}' = R(\sigma')$ .

**Claim 11.** *If  $\hat{d}' > 1.01\hat{d}$ , then  $x \neq y$ , else  $x = y$ .*

*Proof.* The idea is that when  $x = y$ ,  $\hat{d}$  would be really close to  $\hat{d}'$  (up to  $(1 + \epsilon)^2$  because both of them are  $\epsilon$ -approximated) and when  $x \neq y$ , the construction of  $T$  guarantee that  $\hat{d} \geq \hat{d}' + \frac{n}{6}$ . So we can pick an  $\epsilon$  that works for our claim.  $\square$

We can use this method to check every element  $y \in T$  to see if  $y = x$ , and eventually we can recover  $x$  from it. Similar to last section, we can show that  $A$  is an injective function and it implies that  $2^s \geq |T|$  or  $s = \Omega(n)$ .

## 5 Concluding Remarks

- We can use median trick and Chernoff bound to improve the probability of an existing algorithm.
- For distinct elements problem, we can also store the hashes  $h(i)$  approximately. One example is to store the number of leading zeros, and it only cost  $O(\log \log n)$  bits per hash value, and that is the idea behind another algorithm called HyperLogLog.
- For the impossibility results, we can also prove that randomized exact algorithm won't work.